

Causal Treatment Effect Aggregation in Boundary Discontinuity Designs Supplemental Appendix

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Abstract

This supplemental appendix contains theoretical proofs and other technical results.

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SA-1 Preliminary Technical Result

This section is self-contained, and provides the basis for the main analysis in the paper. We employ basic concepts and definitions from geometric measure theory. See Federer [2014] and Folland [2002] for comprehensive reviews; a more accessible source is available at <https://encyclopediaofmath.org/>.

Let $\mathbf{m}$ be the Lebesgue measure, and $\mathfrak{H}^k$ be the k -dimensional Hausdorff measure. For $f : \mathbb{R}^k \rightarrow \mathbb{R}$, let $J_1 f$ denote the Jacobian of f . As in the paper, we assume throughout that $\mathcal{d} : \mathbb{R}^d \times \mathbb{R}^d \rightarrow \mathbb{R}$ is a metric on $\mathbb{R}^d$, and define

$$d(\mathbf{x}, \mathcal{S}) = \inf_{\mathbf{s} \in \mathcal{S}} \mathcal{d}(\mathbf{x}, \mathbf{s}).$$

Assumption SA-1 (Distance).

The following technical lemma gives sufficient conditions for convergence of integrals over a shrinking tubular of a manifold. Let

Lemma SA-1 (Integration over Shrinking Tubular Neighborhoods). *Suppose that $\mathcal{X}$ is a compact subset of $\mathbb{R}^d$ with Hausdorff dimension d , $\mathcal{S} \subseteq \mathcal{X}$ has Hausdorff dimension $d-1$, and the following conditions hold:*

1. *There exists positive constants C_u and C_l such that $C_l \|\mathbf{x}_1 - \mathbf{x}_2\| \leq \mathcal{d}(\mathbf{x}_1, \mathbf{x}_2) \leq C_u \|\mathbf{x}_1 - \mathbf{x}_2\|$ for all $\mathbf{x}_1, \mathbf{x}_2 \in \mathcal{X}$.*
2. *$g : \mathbb{R} \rightarrow \mathbb{R}$ is $\mathbf{m}$ -measurable, continuous over its compact support.*
3. *$m : \mathbb{R}^d \rightarrow \mathbb{R}$ is $\mathbf{m}$ -measurable.*
4. *$J_1 d(\cdot, \mathcal{S}) \neq 0$ $\mathbf{m}$ -almost everywhere on $\mathcal{X}$, and $\int_{\mathcal{X}} \left| \frac{m(\mathbf{x})}{J_1 d(\mathbf{x}, \mathcal{S})} \right| d\mathbf{x} < \infty$.*
5. *$\lim_{\epsilon \downarrow 0} M(\epsilon) = c_{\mathcal{S}} M(0)$ is finite, where $M(r) = \int_{d(\mathbf{x}, \mathcal{S})=r} \frac{m(\mathbf{x})}{J_1 d(\mathbf{x}, \mathcal{S})} d\mathfrak{H}^{d-1}(\mathbf{x})$ for all $r \geq 0$*

Then,

$$\lim_{\epsilon \downarrow 0} \int_{\mathcal{T}(\epsilon)} \frac{1}{\epsilon} g\left(\frac{d(\mathbf{x}, \mathcal{S})}{\epsilon}\right) m(\mathbf{x}) d\mathbf{x} = c_{\mathcal{S}} \int_0^1 g(s) ds \cdot \int_{\mathcal{S}} \frac{m(\mathbf{x})}{J_1 d(\mathbf{x}, \mathcal{S})} d\mathfrak{H}^{d-1}(\mathbf{x}).$$

where

$$\mathcal{T}(\epsilon) = \left\{ \mathbf{x} \in \mathcal{X} : d(\mathbf{x}, \mathcal{S}) \leq \epsilon \right\}, \quad d(\mathbf{x}, \mathcal{S}) = \inf_{\mathbf{y} \in \mathcal{S}} \mathcal{d}(\mathbf{x}, \mathbf{y}).$$

Proof. Note that $d(\cdot, \mathcal{S})$ is C_u -Lipschitz. The level sets

$$\mathcal{L}(\epsilon) = \left\{ \mathbf{x} \in \mathcal{X} : d(\mathbf{x}, \mathcal{S}) = \epsilon \right\}, \quad \epsilon \in [0, \infty)$$

are $(d - 1)$ -dimensional rectifiable sets. Since g is continuous on a compact support, g is also bounded over $\mathcal{X}$. Hence, the function

$$\mathbf{x} \mapsto \epsilon^{-1} g\left(\frac{d(\mathbf{x}, \mathcal{S})}{\epsilon}\right) m(\mathbf{x}) J_1 d(\mathbf{x}, \mathcal{S})^{-1}$$

is $\mathbf{m}$ -summable over $\mathcal{X}$. Using the Coarea formula,

$$\begin{aligned} \int_{\mathcal{T}(\epsilon)} \frac{1}{\epsilon} g\left(\frac{d(\mathbf{x}, \mathcal{S})}{\epsilon}\right) m(\mathbf{x}) d\mathbf{x} &= \int_{\mathcal{T}(\epsilon)} \frac{1}{\epsilon} g\left(\frac{d(\mathbf{x}, \mathcal{S})}{\epsilon}\right) m(\mathbf{x}) \frac{1}{J_1 d(\mathbf{x}, \mathcal{S})} J_1 d(\mathbf{x}, \mathcal{S}) d\mathbf{x} \\ &= \int_0^\epsilon \int_{d(\mathbf{x}, \mathcal{S})=s} \frac{1}{\epsilon} g\left(\frac{s}{\epsilon}\right) \frac{m(\mathbf{x})}{J_1 d(\mathbf{x}, \mathcal{S})} d\mathfrak{H}^{d-1}(\mathbf{x}) ds \\ &= \int_0^\epsilon \frac{1}{\epsilon} g\left(\frac{s}{\epsilon}\right) M(s) ds \\ &= \int_0^1 g(u) M(\epsilon u) du. \end{aligned}$$

Since g is integrable,

$$\lim_{\epsilon \rightarrow 0} \int_0^1 g(u) M(\epsilon u) du = c_{\mathcal{S}} \int_0^1 g(s) ds \cdot M(0).$$

This gives the conclusion. $\square$

SA-2 Setup

This supplemental appendix considers a generalized version of the setup in the main paper: the location variable $\mathbf{X}_i$ is d -dimensional with $d \geq 1$, its support is $\mathcal{X} \subseteq \mathbb{R}^d$, and the boundary region $\mathcal{B}$ is a low dimensional manifold with “effective dimension” $d - 1$. The paper considers the special case $d = 2$, that is, $\mathbf{X}_i$ is bivariate and $\mathcal{B}$ is a one-dimensional (boundary) curve.

We retain the same notation, definitions and assumptions given in the paper, whenever considering a general dimension $d \geq 2$ does not causes confusion. In particular, the natural generalization of Assumption 1 is the following.

Assumption SA-2 (Data Generating Process). *Let $t \in \{0, 1\}$, $p > 0$, and $v \geq 2$.*

- (i) $(Y_1(t), \mathbf{X}_1)^\top, \dots, (Y_n(t), \mathbf{X}_n)^\top$ are independent and identically distributed random vectors.
- (ii) The distribution of $\mathbf{X}_i$ has a Lebesgue density $f(\mathbf{x})$ that is continuous and bounded away from zero on its support compact support $\mathcal{X} \subseteq \mathbb{R}^d$.
- (iii) $\mu_t(\mathbf{x}) = \mathbb{E}[Y_i(t) | \mathbf{X}_i = \mathbf{x}]$ is $(p + 1)$ -times continuously differentiable on $\mathcal{X}$.
- (iv) $\sigma_t^2(\mathbf{x}) = \mathbb{V}[Y_i(t) | \mathbf{X}_i = \mathbf{x}]$ is bounded away from zero and continuous on $\mathcal{X}$.
- (v) $\sup_{\mathbf{x} \in \mathcal{X}} \mathbb{E}[|Y_i(t)|^{2+v} | \mathbf{X}_i = \mathbf{x}] < \infty$.

In addition, we employ the following standard notation through the supplemental appendix.

- (i) *Sample Averages.* $\mathbb{E}_n[g(\mathbf{v}_i)] = \frac{1}{n} \sum_{i=1}^n g(\mathbf{v}_i)$.

- (ii) *Norms.* For a vector $\mathbf{v} \in \mathbb{R}^k$, $\|\mathbf{v}\| = (\sum_{i=1}^k \mathbf{v}_i^2)^{1/2}$ and $\|\mathbf{v}\|_\infty = \max_{1 \leq i \leq k} |\mathbf{v}_i|$. For a matrix $A \in \mathbb{R}^{m \times n}$, $\|A\|_p = \sup_{\|\mathbf{x}\|_p=1} \|A\mathbf{x}\|_p$, $p \in \mathbb{N} \cup \{\infty\}$.
- (iii) *Multi-index Notations.* For a multi-index $\mathbf{u} = (u_1, \dots, u_d) \in \mathbb{N}^d$, denote $|\mathbf{u}| = \sum_{i=1}^d u_i$, $\mathbf{u}! = \prod_{i=1}^d u_i!$. Denote $\mathbf{R}_p(\mathbf{u}) = (1, u_1, \dots, u_d, u_1^2, \dots, u_d^2, \dots, u_1^p, \dots, u_d^p)$, that is, all monomials $u_1^{\alpha_1} \dots u_d^{\alpha_d}$ such that $\alpha_i \in \mathbb{N}$ and $\sum_{i=1}^d \alpha_i \leq p$. Define $\mathbf{e}_{1+\boldsymbol{\nu}}$ to be the $\mathbf{p}_d = \frac{(d+p)!}{d!p!}$ -dimensional vector such that $\mathbf{e}_{1+\boldsymbol{\nu}}^\top \mathbf{R}_p(\mathbf{u}) = \mathbf{u}^\nu$ for all $\mathbf{u} \in \mathbb{R}^d$.
- (iv) *Bounds and Asymptotics.* For reals sequences $a_n = o(b_n)$ if $\limsup \frac{a_n}{b_n} = 0$, and $a_n \lesssim b_n$ or $a_n = O(b_n)$ if there exists some constant C and $N > 0$ such that $n > N$ implies $|a_n| \leq C|b_n|$. For sequences of random variables $A_n = o_{\mathbb{P}}(B_n)$ if $\text{plim}_{n \rightarrow \infty} \frac{A_n}{B_n} = 0$, and $A_n \lesssim_{\mathbb{P}} B_n$ or $A_n = O_{\mathbb{P}}(B_n)$ if $\limsup_{M \rightarrow \infty} \limsup_{n \rightarrow \infty} \mathbb{P}[|\frac{A_n}{B_n}| \geq M] = 0$.

For background textbook references see [van der Vaart and Wellner \[1996\]](#) and [Giné and Nickl \[2016\]](#). For textbook references on geometric measure theory, see [Federer \[2014\]](#) and [Folland \[2002\]](#).

SA-3 Pre-Aggregation Approach

We consider the more general local polynomial estimator described in Remark 1, with $d \geq 2$. Recall the estimator is

$$\hat{\tau}_{\text{pre}} = \mathbf{e}_1^\top \hat{\boldsymbol{\xi}}_1 - \mathbf{e}_1^\top \hat{\boldsymbol{\xi}}_0$$

where

$$\hat{\boldsymbol{\xi}}_t = \arg \min_{\boldsymbol{\xi} \in \mathbb{R}^{p+1}} \sum_{i=1}^n (Y_i - \mathbf{r}_p(D_i)^\top \boldsymbol{\xi})^2 k_h(D_i) \mathbf{1}(T_i = t),$$

for $t \in \{0, 1\}$, and with $\mathbf{e}_1$ the conformable first unit vector, $\mathbf{r}_p(u) = (1, u, u^2, \dots, u^p)^\top$ the usual univariate polynomial basis, $k_h(u) = k(u/h)/h$, $k(\cdot)$ a univariate kernel function, and h the bandwidth.

In addition, define

$$\hat{\theta}_t(r) = \mathbf{r}_p(r)^\top \hat{\boldsymbol{\xi}}_t, \quad r \in \mathbb{R},$$

for $t \in \{0, 1\}$. The induced distance is

$$D_i = (2T_i - 1)d(\mathbf{X}_i, \mathcal{B}), \quad d(\mathbf{x}, \mathcal{B}) = \inf_{\mathbf{b} \in \mathcal{B}} \mathcal{d}(\mathbf{x}, \mathbf{b}),$$

where recall that $\mathbf{X}_i \in \mathbb{R}^d$ and $\mathcal{B} \subset \mathbb{R}^{d-1}$. Thus, the induced conditional expectation based on distance to $\mathcal{B}$ is

$$\theta_t(r) = \mathbb{E}[Y_i | D_i = r] = \mathbb{E}[Y_i | d(\mathbf{X}_i, \mathcal{B}) = |r|, \mathbf{X}_i \in \mathcal{A}_t],$$

for $t \in \{0, 1\}$, and where $r \in \mathcal{J}_t$ with $\mathcal{J}_0 = (-\infty, 0)$ and $\mathcal{J}_1 = [0, \infty)$.

The following assumption is the natural generalization of Assumption 2 in the paper for $d \geq 2$.

Assumption SA-3 (Integral Representation). *Recall the density of $\mathbf{X}_i$, f , from Assumption 1. Suppose the following conditions holds.*

- (i) *There exists positive constants C_u and C_l such that $C_l \|\mathbf{x}_1 - \mathbf{x}_2\| \leq \mathcal{d}(\mathbf{x}_1, \mathbf{x}_2) \leq C_u \|\mathbf{x}_1 - \mathbf{x}_2\|$ for all $\mathbf{x}_1, \mathbf{x}_2 \in \mathcal{X}$.*
- (ii) *$J_1 d(\cdot, \mathcal{B}) \neq 0$ $\mathbf{m}$ -almost everywhere on $\mathcal{X}$, and $\int_{\mathcal{X}} |\frac{f(\mathbf{x})}{J_1 d(\mathbf{x}, \mathcal{B})}| d\mathbf{x} < \infty$.*
- (iii) *$\lim_{\epsilon \downarrow 0} F(s) = F(0)$ is finite, where $F(s) = \int_{d(\mathbf{x}, \mathcal{B})=s} \frac{f(\mathbf{x})}{J_1 d(\mathbf{x}, \mathcal{B})} d\mathfrak{H}^{d-1}(\mathbf{x})$ for all $s \geq 0$.*
- (iv) *θ_t is continuous at 0 for $t = 0, 1$.*

For each $t \in \{0, 1\}$, the best mean square approximation based on $\mathbf{r}_p(D_i)$ is

$$\theta_t^*(D_i) = \mathbf{r}_p(D_i)^\top \boldsymbol{\gamma}_t^*$$

with

$$\boldsymbol{\gamma}_t^* = \arg \min_{\boldsymbol{\gamma} \in \mathbb{R}^{p+1}} \mathbb{E} \left[\left(Y_i - \mathbf{r}_p(D_i)^\top \boldsymbol{\gamma} \right)^2 k_h(D_i) \mathbf{1}_{\mathcal{J}_t}(D_i) \right].$$

Thus, the estimation error decomposes into *linear error*, *approximation error*, and *non-linear error*:

$$\begin{aligned} \hat{\theta}_t(0) - \theta_t(0) &= \mathbf{e}_1^\top \hat{\boldsymbol{\Psi}}_t^{-1} \mathbb{E}_n \left[\mathbf{r}_p \left(\frac{D_i}{h} \right) k_h(D_i) Y_i \right] - \theta_t(0) \\ &= \mathbf{e}_1^\top \hat{\boldsymbol{\Psi}}_t^{-1} \mathbb{E}_n \left[\mathbf{r}_p \left(\frac{D_i}{h} \right) k_h(D_i) (Y_i - \theta_t^*(D_i)) \right] + \theta_t^*(0) - \theta_t(0) \\ &= \underbrace{\mathbf{e}_1^\top \boldsymbol{\Psi}_t^{-1} \mathbf{O}_t}_{\text{linear error}} + \underbrace{\theta_t^*(0) - \theta_t(0)}_{\text{approximation error}} + \underbrace{\mathbf{e}_1^\top (\hat{\boldsymbol{\Psi}}_t^{-1} - \boldsymbol{\Psi}_t^{-1}) \mathbf{O}_t}_{\text{non-linear error}}, \end{aligned} \tag{SA-1}$$

for $t \in \{0, 1\}$, where

$$\begin{aligned} \hat{\boldsymbol{\Psi}}_t &= \mathbb{E}_n \left[\mathbf{r}_p \left(\frac{D_i}{h} \right) \mathbf{r}_p \left(\frac{D_i}{h} \right)^\top k_h(D_i) \mathbf{1}_{\mathcal{J}_t}(D_i) \right], \\ \boldsymbol{\Psi}_t &= \mathbb{E} \left[\mathbf{r}_p \left(\frac{D_i}{h} \right) \mathbf{r}_p \left(\frac{D_i}{h} \right)^\top k_h(D_i) \mathbf{1}_{\mathcal{J}_t}(D_i) \right], \\ \mathbf{O}_t &= \mathbb{E}_n \left[\mathbf{r}_p \left(\frac{D_i}{h} \right) k_h(D_i) (Y_i - \theta_t^*(D_i)) \mathbf{1}_{\mathcal{J}_t}(D_i) \right], \end{aligned}$$

and the misspecification bias is

$$\mathfrak{B}_{n,t} = \theta_t^*(0) - \theta_t(0). \tag{SA-2}$$

In the text, $\mathfrak{B}_n = \mathfrak{B}_{n,1} - \mathfrak{B}_{n,0}$. In addition, define

$$\begin{aligned}\hat{\Upsilon}_t &= h\mathbb{E}_n\left[\mathbf{r}_p\left(\frac{D_i}{h}\right)\mathbf{r}_p\left(\frac{D_i}{h}\right)^\top k_h(D_i)^2(Y_i - \hat{\theta}_t(D_i))^2\mathbf{1}_{\mathcal{J}_t}(D_i)\right], \\ \Upsilon_t &= h\mathbb{E}\left[\mathbf{r}_p\left(\frac{D_i}{h}\right)\mathbf{r}_p\left(\frac{D_i}{h}\right)^\top k_h(D_i)^2(Y_i - \theta_t^*(D_i))^2\mathbf{1}_{\mathcal{J}_t}(D_i)\right], \\ \hat{\Xi}_t &= \frac{1}{nh}\mathbf{e}_1^\top \hat{\Psi}_t^{-1} \hat{\Upsilon}_t \hat{\Psi}_t^{-1} \mathbf{e}_1, \quad \hat{\Xi} = \hat{\Xi}_0 + \hat{\Xi}_1, \\ \Xi_t &= \frac{1}{nh}\mathbf{e}_1^\top \Psi_t^{-1} \Upsilon_t \Psi_t^{-1} \mathbf{e}_1, \quad \Xi = \Xi_0 + \Xi_1,\end{aligned}$$

and

$$\rho_t(r) = \mathbb{E}[(Y_i - \theta_t^*(D_i))^2 | D_i = r] = \mathbb{E}[Y_i^2 | D_i = r] - 2\theta_t(r)\mathbf{r}_p(r)^\top \boldsymbol{\gamma}_t^* + (\mathbf{r}_p(r)^\top \boldsymbol{\gamma}_t^*)^2.$$

for $t \in \{0, 1\}$.

SA-3.1 Preliminary Lemmas

Lemma SA-2 (Expected Gram Matrix). *Suppose Assumptions SA-2 and SA-3 hold. If $h \rightarrow 0$, then*

$$\left\| \Psi_t - \int_{\mathcal{B}} \frac{f(\mathbf{x})}{J_1 d(\mathbf{x}, \mathcal{B})} d\mathbf{x} \cdot \mathbf{\Gamma} \right\| = o(1), \quad \mathbf{\Gamma} = \int_0^1 \mathbf{r}_p(u) \mathbf{r}_p(u)^\top k(u) du,$$

for $t \in \{0, 1\}$. This implies $\|\Psi_t\| \lesssim 1$ and $\|\Psi_t^{-1}\| \lesssim 1$.

Proof. The typical entry of Ψ_t has the form

$$\psi_{n,t} = \mathbb{E}[(D_i/h)^v k_h(D_i) \mathbf{1}_{\mathcal{J}_t}(D_i)],$$

for some $v \in \mathbb{N}$. By Lemma SA-1, as $h \rightarrow 0$,

$$\psi_{n,t} = \int_{\mathcal{R}_t(h)} \left(\frac{d(\mathbf{x}, \mathcal{B})}{h}\right)^v \frac{1}{h} k\left(\frac{d(\mathbf{x}, \mathcal{B})}{h}\right) f(\mathbf{x}) d\mathbf{x} = \int_0^1 s^v k(s) ds \cdot \int_{\mathcal{B}} \frac{f(\mathbf{x})}{J_1 d(\mathbf{x}, \mathcal{B})} d\mathbf{x} + o(1).$$

Let $\mathbf{p}_p = \frac{(d+p)!}{d!p!}$. Since ψ_t is $\mathbf{p}_p \times \mathbf{p}_p$ dimensional, the conclusion follows from a union bound. By Weyl's theorem, $|\lambda_{\min}(\Psi_t) - \lambda_{\min}(\mathbf{\Gamma})| = o(1)$ and $|\lambda_{\max}(\Psi_t) - \lambda_{\max}(\mathbf{\Gamma})| = o(1)$. Hence $\|\Psi_t\| \lesssim 1$ and $\|\Psi_t^{-1}\| \lesssim 1$. $\square$

Lemma SA-3 (Estimated Gram Matrix). *Suppose Assumptions SA-2 and SA-3 hold. If $h \rightarrow 0$ and $nh \rightarrow \infty$, then*

$$\|\hat{\Psi}_t - \Psi_t\| \lesssim_{\mathbb{P}} (nh)^{-1/2},$$

for $t \in \{0, 1\}$. This implies $\|\hat{\Psi}_t\| \lesssim_{\mathbb{P}} 1$ and $\|\hat{\Psi}_t^{-1}\| \lesssim_{\mathbb{P}} 1$.

Proof. The typical entry of $\widehat{\Psi}_t$ takes the form

$$\psi_{n,t} = \frac{1}{n} \sum_{i=1}^n \left(\frac{D_i}{h} \right)^v \frac{1}{h} k \left(\frac{D_i}{h} \right) \mathbf{1}_{\mathcal{J}_t}(D_i).$$

By Lemma SA-1,

$$\mathbb{E} \left[\left(\frac{D_i}{h} \right)^{2v} \frac{1}{h} k \left(\frac{D_i}{h} \right)^2 \mathbf{1}_{\mathcal{J}_t}(D_i) \right] = \int_0^1 s^{2v} k(s)^2 ds \cdot \int_{\mathcal{B}} \frac{f(\mathbf{x})}{J_1 d(\mathbf{x}, \mathcal{B})} d\mathbf{x}.$$

Hence,

$$\begin{aligned} \mathbb{V}[\psi_{n,t}] &= \frac{1}{nh^2} \mathbb{V} \left[\left(\frac{D_i}{h} \right)^v k \left(\frac{D_i}{h} \right) \mathbf{1}_{\mathcal{J}_t}(D_i) \right] \\ &= \frac{1}{nh} \mathbb{E} \left[\left(\frac{D_i}{h} \right)^{2v} \frac{1}{h} k \left(\frac{D_i}{h} \right)^2 \mathbf{1}_{\mathcal{J}_t}(D_i) \right] - \frac{1}{n} \mathbb{E} \left[\left(\frac{D_i}{h} \right)^v \frac{1}{h} k \left(\frac{D_i}{h} \right) \mathbf{1}_{\mathcal{J}_t}(D_i) \right]^2 \\ &= \frac{1}{nh} \int_0^1 s^{2v} k(s)^2 ds \cdot \int_{\mathcal{B}} \frac{f(\mathbf{x})}{J_1 d(\mathbf{x}, \mathcal{B})} d\mathbf{x} (1 + o(1)) + \frac{1}{n} \int_0^1 s^v k(s) ds \cdot \int_{\mathcal{B}} \frac{f(\mathbf{x})}{J_1 d(\mathbf{x}, \mathcal{B})} d\mathbf{x} (1 + o(1)) \\ &= \frac{1}{nh} \int_0^1 s^{2v} k(s)^2 ds \cdot \int_{\mathcal{B}} \frac{f(\mathbf{x})}{J_1 d(\mathbf{x}, \mathcal{B})} d\mathbf{x} (1 + o(1)). \end{aligned}$$

Since Ψ_t is finite dimensional, $\|\widehat{\Psi}_t - \Psi_t\| \lesssim_{\mathbb{P}} (nh)^{-1/2}$. The conclusion then follows from Lemma SA-3 and Weyl's theorem. $\square$

Lemma SA-4 (Stochastic Linear Approximation). *Suppose Assumptions SA-2 and SA-3 hold. If $nh \rightarrow \infty$, then*

$$\begin{aligned} \|\mathbf{O}_t\| &\lesssim_{\mathbb{P}} (nh)^{-1/2}, \\ \left| \mathbf{e}_1^\top \Psi_t^{-1} \mathbf{O}_t \right| &\lesssim_{\mathbb{P}} (nh)^{-1/2}, \\ \left| \mathbf{e}_1^\top (\widehat{\Psi}_t^{-1} - \Psi_t^{-1}) \mathbf{O}_t \right| &\lesssim_{\mathbb{P}} (nh)^{-1}, \end{aligned}$$

for $t \in \{0, 1\}$.

Proof. By the definition of the best linear approximation θ_t^* , we have $\mathbb{E}[\mathbf{O}_t] = 0$. A typical entry of $\mathbf{O}_t$ has the form

$$o_t = \mathbb{E}_n \left[\left(\frac{D_i}{h} \right)^v k_h(D_i) (Y_i - \theta_t^*(D_i)) \mathbf{1}_{\mathcal{J}_t}(D_i) \right]$$

for $v \in \mathbb{N}$. Since $\sup_{r \in \mathbb{R}} \mathbf{1}(k(r/h) > 0) \rho_t(r) \lesssim 1$, by Lemma SA-1,

$$\begin{aligned} \mathbb{V}[o_t] &= \frac{1}{nh} \mathbb{E} \left[\left(\frac{D_i}{h} \right)^{2v} \frac{1}{h} k \left(\frac{D_i}{h} \right)^2 \rho_t(D_i) \mathbf{1}_{\mathcal{J}_t}(D_i) \right] \\ &\lesssim \frac{1}{nh} \mathbb{E} \left[\left(\frac{D_i}{h} \right)^{2v} \frac{1}{h} k \left(\frac{D_i}{h} \right)^2 \mathbf{1}_{\mathcal{J}_t}(D_i) \right] \end{aligned}$$

$$\lesssim \frac{1}{nh} \int_0^1 s^{2v} k(s)^2 ds.$$

Since $\mathbf{O}_t$ is finite dimensional, $\|\mathbf{O}_t\| \lesssim (nh)^{-1/2}$. The other results follow from Lemma SA-2 and Lemma SA-3. $\square$

Lemma SA-5 (Expected Meat Matrix). *Suppose Assumptions SA-2 and SA-3 hold. If $h \rightarrow 0$, then*

$$\left\| \Upsilon_t - \iota_t \int_{\mathcal{B}} \frac{f(\mathbf{x})}{J_1 d(\mathbf{x}, \mathcal{B})} d\mathbf{x} \cdot \Sigma \right\| = o(1), \quad \Sigma = \int_0^1 \mathbf{r}_p(u) \mathbf{r}_p(u)^\top k(u)^2 du,$$

and

$$\iota_t = \int_{\mathcal{B}} \sigma_t^2(\mathbf{b}) \frac{f(\mathbf{b})}{J_1 d(\mathbf{b}, \mathcal{B})} d\mathfrak{H}^{d-1}(\mathbf{b}) + \int_{\mathcal{B}} \left(\mu_t(\mathbf{b}) - \frac{\int_{\mathcal{B}} \mu_t(\mathbf{x}) \frac{f(\mathbf{x})}{J_1 d(\mathbf{x}, \mathcal{B})} d\mathbf{x}}{\int_{\mathcal{B}} \frac{f(\mathbf{u})}{J_1 d(\mathbf{u}, \mathcal{B})} d\mathbf{u}} \right)^2 \frac{f(\mathbf{b})}{J_1 d(\mathbf{b}, \mathcal{B})} d\mathfrak{H}^{d-1}(\mathbf{b}),$$

for $t \in \{0, 1\}$.

Proof. Consider $s_t(\mathbf{x}) = \mathbb{E}[Y_i(t)^2 | \mathbf{X}_i = \mathbf{x}] = \mu_t(\mathbf{x})^2 + \sigma_t(\mathbf{x})^2$, $\mathbf{x} \in \mathcal{X}$. A typical element of Υ_t has the form

$$\begin{aligned} v_t &= h \mathbb{E} \left[\left(\frac{D_i}{h} \right)^v k_h(D_i)^2 (Y_i - \theta_t^*(D_i))^2 \mathbf{1}_{\mathcal{J}_t}(D_i) \right] \\ &= \int_{\mathcal{R}_t(h)} \left(\frac{d(\mathbf{x}, \mathcal{B})}{h} \right)^v \frac{1}{h} k \left(\frac{d(\mathbf{x}, \mathcal{B})}{h} \right)^2 (s_t(\mathbf{x}) - 2\mu_t(\mathbf{x})\theta_t^*(d(\mathbf{x}, \mathcal{B})) + \theta_t^*(d(\mathbf{x}, \mathcal{B}))^2) f(\mathbf{x}) d\mathbf{x} \\ &= \int_0^1 s^v k(s)^2 ds \left[\int_{\mathcal{B}} \frac{s_t(\mathbf{x}) f(\mathbf{x})}{J_1 d(\mathbf{x}, \mathcal{B})} d\mathbf{x} - 2 \int_{\mathcal{B}} \frac{\mu_t(\mathbf{x}) f(\mathbf{x})}{J_1 d(\mathbf{x}, \mathcal{B})} d\mathbf{x} \cdot \theta_t^*(0) + \int_{\mathcal{B}} \frac{f(\mathbf{x})}{J_1 d(\mathbf{x}, \mathcal{B})} d\mathbf{x} \theta_t^*(0)^2 \right] + o(1), \end{aligned}$$

where in the last line we have used Lemma SA-1, and $\theta_t^*(r) = \mathbf{R}_p(r)^\top \boldsymbol{\gamma}_t^*$ is continuous in r . Moreover,

$$\theta_t^*(0) = \mathbf{e}_1^\top \boldsymbol{\Psi}_t^{-1} \mathbb{E} \left[\mathbf{r}_p \left(\frac{D_i}{h} \right) k_h(D_i) Y_i \mathbf{1}_{\mathcal{J}_t}(D_i) \right].$$

Using Lemma SA-1 again, we have

$$\begin{aligned} \mathbb{E} \left[\mathbf{r}_p \left(\frac{D_i}{h} \right) k_h(D_i) Y_i \mathbf{1}_{\mathcal{J}_t}(D_i) \right] &= \int_{\mathcal{R}_t(h)} \mathbf{r}_p \left(\frac{d(\mathbf{x}, \mathcal{B})}{h} \right) k_h(d(\mathbf{x}, \mathcal{B})) \mu_t(\mathbf{x}) \mathbf{1}_{\mathcal{J}_t}(D_i) d\mathfrak{H}^{d-1}(\mathbf{x}) \\ &= \boldsymbol{\Lambda} \int_{\mathcal{B}} \frac{\mu_t(\mathbf{x}) f(\mathbf{x})}{J_1 d(\mathbf{x}, \mathcal{B})} d\mathfrak{H}^{d-1}(\mathbf{x}), \end{aligned}$$

where $\boldsymbol{\Lambda} = \int_0^1 \mathbf{r}_p(u) k(u) du$. Together with Lemma SA-2, we have

$$\theta_t^*(0) = \mathbf{e}_1^\top \boldsymbol{\Gamma}^{-1} \boldsymbol{\Lambda} \frac{\int_{\mathcal{B}} \mu_t(\mathbf{x}) f(\mathbf{x}) (J_1 d(\mathbf{x}, \mathcal{B}))^{-1} d\mathfrak{H}^{d-1}(\mathbf{x})}{\int_{\mathcal{B}} f(\mathbf{x}) (J_1 d(\mathbf{x}, \mathcal{B}))^{-1} d\mathfrak{H}^{d-1}(\mathbf{x})},$$

where the definition of $\mathbf{\Gamma}$ from Lemma SA-2 implies $\mathbf{e}_1^\top \mathbf{\Gamma}^{-1} \mathbf{\Lambda} = 1$. Finite dimensionality and the union bound over the entries give the result. $\square$

Lemma SA-6 (Covariance). *Suppose Assumptions SA-2 and SA-3 hold. If $h \rightarrow 0$ and $nh \rightarrow \infty$, then*

$$\begin{aligned} \left\| \hat{\mathbf{\Upsilon}}_t - \mathbf{\Upsilon}_t \right\| &\lesssim_{\mathbb{P}} (nh)^{-1/2}, \\ nh |\hat{\Xi}_t - \Xi_t| &\lesssim_{\mathbb{P}} (nh)^{-1/2}, \end{aligned}$$

for $t \in \{0, 1\}$. This implies $1 \lesssim_{\mathbb{P}} \lambda_{\min}(\hat{\mathbf{\Upsilon}}_t) \lesssim_{\mathbb{P}} 1$ and $(nh)^{-1} \lesssim_{\mathbb{P}} \hat{\Xi}_t \lesssim_{\mathbb{P}} (nh)^{-1}$.

Proof. Denote $\eta_{i,t} = Y_i - \theta_t^*(D_i)$ and $\xi_{i,t} = \theta_t^*(D_i) - \hat{\theta}_t(D_i)$. Then

$$\hat{\mathbf{\Upsilon}}_t = \mathbb{E}_n \left[\mathbf{r}_p \left(\frac{D_i}{h} \right) \mathbf{r}_p \left(\frac{D_i}{h} \right)^\top h k_h(D_i) k_h(D_i) (\eta_{i,t} + \xi_{i,t})^2 \mathbf{1}_{\mathcal{J}_t}(D_i) \right],$$

and we decompose the error into

$$\begin{aligned} \hat{\mathbf{\Upsilon}}_t - \mathbf{\Upsilon}_t &= \Delta_{1,t} + \Delta_{2,t} + \Delta_{3,t}, \\ \Delta_{1,t} &= \mathbb{E}_n \left[\mathbf{r}_p \left(\frac{D_i}{h} \right) \mathbf{r}_p \left(\frac{D_i}{h} \right)^\top h k_h(D_i) k_h(D_i) \xi_{i,t}^2 \mathbf{1}_{\mathcal{J}_t}(D_i) \right], \\ \Delta_{2,t} &= 2 \mathbb{E}_n \left[\mathbf{r}_p \left(\frac{D_i}{h} \right) \mathbf{r}_p \left(\frac{D_i}{h} \right)^\top h k_h(D_i) k_h(D_i) \eta_{i,t} \xi_{i,t} \mathbf{1}_{\mathcal{J}_t}(D_i) \right], \\ \Delta_{3,t} &= \mathbb{E}_n \left[\mathbf{r}_p \left(\frac{D_i}{h} \right) \mathbf{r}_p \left(\frac{D_i}{h} \right)^\top h k_h(D_i) k_h(D_i) \eta_{i,t}^2 \mathbf{1}_{\mathcal{J}_t}(D_i) \right] \\ &\quad - \mathbb{E} \left[\mathbf{r}_p \left(\frac{D_i}{h} \right) \mathbf{r}_p \left(\frac{D_i}{h} \right)^\top h k_h(D_i) k_h(D_i) \eta_{i,t}^2 \mathbf{1}_{\mathcal{J}_t}(D_i) \right]. \end{aligned}$$

$k_h(D_i) \neq 0$ implies $\|\mathbf{r}_p(D_i/h)\|_2 \lesssim 1$. Define

$$\begin{aligned} \mathbf{U}_t &= \mathbb{E}_n \left[\mathbf{r}_p \left(\frac{D_i}{h} \right) k_h(D_i) \theta_t^*(D_i) \mathbf{1}_{\mathcal{J}_t}(D_i) \right], \\ \mathbf{H} &= \text{diag}(1, h, \dots, h^p). \end{aligned}$$

Hence, by Lemma SA-3 and SA-4,

$$\begin{aligned} &\max_{t \in \{0,1\}} \max_{1 \leq i \leq n} |\xi_{i,t}| \\ &= \max_{t \in \{0,1\}} \max_{1 \leq i \leq n} |\mathbf{r}_p(D_i)^\top (\hat{\gamma}_t - \gamma_t^*)| \mathbf{1}(k_h(D_i) \geq 0) \\ &= \max_{t \in \{0,1\}} \max_{1 \leq i \leq n} |\mathbf{r}_p(D_i)^\top \mathbf{H}^{-1} (\hat{\Psi}_t^{-1} \mathbf{O}_t + (\hat{\Psi}_t^{-1} - \Psi_t^{-1}) \mathbf{U}_t)| \mathbf{1}(k_h(D_i) \geq 0) \\ &\leq \max_{t \in \{0,1\}} \left\| \hat{\Psi}_t^{-1} \mathbf{O}_t \right\|_2 + \max_{t \in \{0,1\}} \left\| (\hat{\Psi}_t^{-1} - \Psi_t^{-1}) \mathbf{U}_t \right\|_2 \end{aligned}$$

$$\lesssim_{\mathbb{P}} (nh)^{-1/2},$$

Assume $nh \rightarrow \infty$, similar maximal inequality as in the proof of Lemma SA-3 shows

$$\begin{aligned} \max_{t=0,1} \|\Delta_{1,t}\| &\lesssim_{\mathbb{P}} \max_{t \in \{0,1\}} \max_{1 \leq i \leq n} |\xi_{i,t}|^2 \lesssim_{\mathbb{P}} (nh)^{-1}, \\ \max_{t=0,1} \|\Delta_{2,t}\| &\lesssim_{\mathbb{P}} \max_{t \in \{0,1\}} \max_{1 \leq i \leq n} |\xi_{i,t}| \lesssim_{\mathbb{P}} (nh)^{-1/2}. \end{aligned} \quad (\text{SA-3})$$

For $\Delta_{3,t}$, notice that a typical entry of has the form

$$\mathbf{g}_n - \mathbb{E}[\mathbf{g}_n],$$

where

$$\mathbf{g}_n = \frac{1}{n} \sum_{i=1}^n \left(\frac{D_i}{h} \right)^v \frac{1}{h} k \left(\frac{D_i}{h} \right)^2 (Y_i - \theta_t^*(D_i))^2 \mathbf{1}_{\mathcal{J}_t}(D_i).$$

Since we have assumed $\sup_{\mathbf{x} \in \mathcal{X}} \mathbb{E}[Y_t^4 | \mathbf{X}_i = \mathbf{x}] < \infty$, and by Jensen's inequality $\mathbb{E}[\theta_t^*(D_i)^4] < \infty$, a similar argument as for Lemma SA-3 implies

$$\mathbb{V}[\mathbf{g}_n] \lesssim (nh)^{-1},$$

and hence

$$\max_{t=0,1} \|\Delta_{3,t}\| \lesssim_{\mathbb{P}} (nh)^{-1/2}. \quad (\text{SA-4})$$

Putting together Equation (SA-3) and (SA-4), we get $\|\hat{\Upsilon}_t - \Upsilon_t\| \lesssim_{\mathbb{P}} (nh)^{-1/2}$. By Lemma SA-5 and Weyl's theorem, we have $1 \lesssim_{\mathbb{P}} \lambda_{\min}(\hat{\Upsilon}_t) \lesssim_{\mathbb{P}} 1$. Using Lemma SA-3 in addition, we can show

$$nh|\hat{\Xi}_t - \Xi_t| \lesssim_{\mathbb{P}} |\mathbf{e}_1^\top \hat{\Psi}_t^{-1} \hat{\Upsilon}_t \hat{\Psi}_t^{-1} \mathbf{e}_1 - \mathbf{e}_1^\top \Psi_t^{-1} \Upsilon_t \Psi_t^{-1} \mathbf{e}_1| \lesssim_{\mathbb{P}} (nh)^{-1/2}.$$

Lemma SA-2 and SA-5 imply $(nh)^{-1} \lesssim \Xi_t \lesssim (nh)^{-1}$, hence $(nh)^{-1} \lesssim_{\mathbb{P}} \hat{\Xi}_t \lesssim_{\mathbb{P}} (nh)^{-1}$. $\square$

SA-3.2 Mean Square Error Approximation

Theorem SA-1 (MSE Expansion). *Suppose Assumptions SA-2 and SA-3 hold. If $h \rightarrow 0$ and $nh \rightarrow \infty$, then*

$$\mathbb{E} \left[\left(\sum_{t=0,1} (-1)^{t+1} (\mathbf{e}_1^\top \Psi_t^{-1} \mathbf{O}_t + \theta_t^*(0) - \theta_t(0)) \right)^2 \right] = \mathfrak{B}_n^2 + \Xi + O_{\mathbb{P}}((nh)^{-3/2}).$$

Proof. By definition of Υ_t we have,

$$\begin{aligned}\mathbb{E}[\mathbf{O}_t \mathbf{O}_t^\top] &= \frac{1}{nh} \mathbb{E} \left[\mathbf{r}_p \left(\frac{D_i}{h} \right) \mathbf{r}_p \left(\frac{D_i}{h} \right)^\top h k_h(D_i)^2 (Y_i - \theta_t^*(D_i))^2 \mathbf{1}_{\mathcal{J}_t}(D_i) \right] \\ &= \frac{1}{nh} \Upsilon_t.\end{aligned}$$

Hence

$$\begin{aligned}\mathbb{E} &\left[\left(\sum_{t=0,1} (-1)^{t+1} (\mathbf{e}_1^\top \Psi_t^{-1} \mathbf{O}_t + (\theta_t^*(0) - \theta_t(0))) \right)^2 \right] \\ &= 2\mathbb{E} \left[\left(\sum_{t=0,1} (-1)^{t+1} (\mathbf{e}_1^\top \Psi_t^{-1} \mathbf{O}_t) \right) \left(\sum_{t=0,1} (-1)^{t+1} (\theta_t^*(0) - \theta_t(0)) \right) \right] \\ &\quad + \mathbb{E} \left[\left(\sum_{t=0,1} (-1)^{t+1} (\mathbf{e}_1^\top \Psi_t^{-1} \mathbf{O}_t) \right)^2 \right] + \left(\sum_{t=0,1} (-1)^{t+1} (\theta_t^*(0) - \theta_t(0)) \right)^2 \\ &= \Xi_0 + \Xi_1 + \left(\sum_{t=0,1} (-1)^{t+1} (\theta_t^*(0) - \theta_t(0)) \right)^2,\end{aligned}$$

where in the third line we have used $\mathbb{E}[\mathbf{O}_t] = \mathbf{0}$ for $t = 0, 1$, and the independence between $\mathbf{O}_0$ and $\mathbf{O}_1$. The conclusion follows. $\square$

Corollary SA-1 (Convergence Rate). *Suppose Assumption 1, SA-3 hold, and $nh \rightarrow \infty$. Then*

$$|\hat{\tau}_{\text{pre}} - \tau_f| \lesssim_{\mathbb{P}} (nh)^{-1/2} + |\mathfrak{B}_n|.$$

Proof. The conclusion follows from Lemma SA-2, SA-5 and Theorem SA-1. $\square$

SA-3.3 Central Limit Theorem

The feasible t-statistics by

$$\hat{\mathbf{T}}_{\text{pre}} = \frac{\hat{\tau}_{\text{pre}} - \tau_f}{\sqrt{\hat{\Xi}}}.$$

Theorem SA-2 (Asymptotic Normality). *Suppose Assumptions SA-2 and SA-3 hold. If $nh \rightarrow \infty$ and $\sqrt{nh}|\mathfrak{B}_n| \rightarrow 0$, then*

$$\sup_{u \in \mathbb{R}} |\mathbb{P}(\hat{\tau}_{\text{pre}} \leq u) - \Phi(u)| = o(1).$$

Proof. Define the stochastic linearization of $\hat{\mathbf{T}}_{\text{pre}}$ to be $\bar{\mathbf{T}}_{\text{pre}} = \Xi^{-1/2} \mathbf{e}_1^\top \Psi_t^{-1} \mathbf{O}_t$. First, we bound

the stochastic linearization error. Using the decomposition (SA-1) and convergence of $\widehat{\Xi}$,

$$\begin{aligned}\widehat{\mathbf{T}}_{\text{pre}} - \overline{\mathbf{T}}_{\text{pre}} &= \widehat{\Xi}^{-1/2} \left(\sum_{t \in \{0,1\}} (-1)^{\frac{t+1}{2}} (\widehat{\theta}_t(0) - \theta_t(0)) \right) - \Xi^{-1/2} \left(\sum_{t \in \{0,1\}} (-1)^{\frac{t+1}{2}} \mathbf{e}_1^\top \boldsymbol{\Psi}_t^{-1} \mathbf{O}_t \right) \\ &= \widehat{\Xi}^{-1/2} \left(\sum_{t \in \{0,1\}} (-1)^{\frac{t+1}{2}} (\widehat{\theta}_t(0) - \theta_t(0)) - \sum_{t \in \{0,1\}} (-1)^{\frac{t+1}{2}} \mathbf{e}_1^\top \boldsymbol{\Psi}_t^{-1} \mathbf{O}_t \right) \quad (= \Delta_1) \\ &\quad + (\widehat{\Xi}^{-1/2} - \Xi^{-1/2}) \sum_{t \in \{0,1\}} (-1)^{\frac{t+1}{2}} \mathbf{e}_1^\top \boldsymbol{\Psi}_t^{-1} \mathbf{O}_t \quad (= \Delta_2)\end{aligned}$$

By Lemma SA-3 and SA-4, and the decomposition Equation (SA-1),

$$\sup_{\mathbf{x} \in \mathcal{X}} \left| \sum_{t \in \{0,1\}} (-1)^{\frac{t+1}{2}} (\widehat{\theta}_t(0) - \theta_t(0)) - \sum_{t \in \{0,1\}} (-1)^{\frac{t+1}{2}} \mathbf{e}_1^\top \boldsymbol{\Psi}_t^{-1} \mathbf{O}_t \right| \lesssim_{\mathbb{P}} (nh)^{-1} + |\mathfrak{B}_n|.$$

Together with Lemma SA-6,

$$|\Delta_1| \lesssim_{\mathbb{P}} (nh)^{-1/2} + \sqrt{nh} |\mathfrak{B}_n|. \quad (\text{SA-5})$$

By Lemma SA-3, Lemma SA-4 and Lemma SA-6, and assume $nh \rightarrow \infty$, then

$$|\Delta_2| = \sum_{t \in \{0,1\}} \left| \mathbf{e}_1^\top \boldsymbol{\Psi}_t^{-1} \mathbf{O}_t \left(\Xi^{-1/2} - \widehat{\Xi}^{-1/2} \right) \right| \lesssim_{\mathbb{P}} (nh)^{-1/2}. \quad (\text{SA-6})$$

Putting together Equations (SA-5), (SA-6) gives

$$|\widehat{\mathbf{T}}_{\text{pre}} - \overline{\mathbf{T}}_{\text{pre}}| \lesssim_{\mathbb{P}} (nh)^{-1/2} + \sqrt{nh} |\mathfrak{B}_n|.$$

Next, we consider the convergence of $\overline{\mathbf{T}}_{\text{pre}}$. Notice that if we define

$$Z_{n,i} = \frac{1}{n} \Xi^{-1/2} \mathbf{e}_1^\top \boldsymbol{\Psi}_t^{-1} \mathbf{r}_p \left(\frac{D_i}{h} \right) k_h(D_i) (Y_i - \theta_t^*(D_i)) \mathbf{1}_{\mathcal{J}_t}(D_i),$$

then $\overline{\mathbf{T}}_{\text{pre}} = \sum_{i=1}^n Z_{n,i}$. Moreover, $\mathbb{E}[Z_{n,i}] = 0$ and $\mathbb{V}[Z_{n,i}] = n^{-1}$. By Berry-Essen Theorem,

$$\begin{aligned}\sup_{u \in \mathbb{R}} |\mathbb{P}(\overline{\mathbf{T}}_{\text{pre}} \leq u) - \Phi(u)| &\lesssim \sum_{i=1}^n \mathbb{E}[|Z_{n,i}|^3] \\ &= \sum_{i=1}^n n^{-3} \Xi^{-3/2} \mathbb{E} \left[\left| \mathbf{e}_1^\top \boldsymbol{\Psi}_t^{-1} \mathbf{r}_p \left(\frac{D_i}{h} \right) k_h(D_i) \mathbf{1}_{\mathcal{J}_t}(D_i) (Y_i - \theta_t^*(D_i)) \right|^3 \right] \\ &\lesssim n^{-2} \Xi^{-3/2} \mathbb{E}[|k_h(D_i)(Y_i - \theta_t^*(D_i))|^3] \\ &\lesssim n^{-2} \Xi^{-3/2} \mathbb{E}[|k_h(D_i)|^3 (\mathbb{E}[|Y_i|^3 | \mathbf{X}_i] + |\theta_t^*(D_i)|^3)] \\ &\lesssim (nh)^{-1/2},\end{aligned}$$

where in the third line we used $\|\mathbf{r}_p \left(\frac{D_i}{h} \right) k_h(D_i)\| \lesssim 1$ holds almost surely, in last line we have use Lemma SA-1 with Assumptions 1 and SA-3 to get $\mathbb{E}[|k_h(D_i)|^3(\mathbb{E}[|Y_i|^3|\mathbf{X}_i] + |\theta_t^*(D_i)|^3)] \lesssim h^{-2}$, and the fact that $\Xi \gtrsim (nh)^{-1/2}$ from Lemma SA-6. $\square$

SA-3.4 Approximation Error

Lemma SA-7 (Approximation Error). *Suppose Assumptions SA-2 and SA-3 hold. If θ_t is $(s+1)$ -times continuously differentiable at 0, and $h \rightarrow 0$, then*

$$\mathfrak{B}_{n,t} \lesssim h^{\min\{s,p\}+1}$$

for $t \in \{0, 1\}$.

Proof. Consider

$$\mathbf{S}_t = \mathbb{E} \left[\mathbf{r}_p \left(\frac{D_i}{h} \right) k_h(D_i) Y_i \mathbf{1}_{\mathcal{J}_t}(D_i) \right] = \mathbb{E} \left[\mathbf{r}_p \left(\frac{D_i}{h} \right) k_h(D_i) \theta_t(D_i) \mathbf{1}_{\mathcal{J}_t}(D_i) \right].$$

Define $\eta_t(h)$ to be the p -dimensional vector that is $(\theta_t(0), h\theta_t^{(1)}(0), \dots, h^p\theta_t^{(p)}(0))$ if $p \leq s$, and $(\theta_t(0), h\theta_t^{(1)}(0), \dots, h^s\theta_t^{(s)}(0), 0, \dots, 0)$ otherwise.

$$\begin{aligned} |\mathfrak{B}_{n,t}| &= \left| \mathbf{e}_1^\top \boldsymbol{\Psi}_t^{-1} \mathbf{S}_t - \theta_t(0) \right| \\ &= \left| \mathbf{e}_1^\top \boldsymbol{\Psi}_t^{-1} \mathbb{E} \left[\mathbf{r}_p \left(\frac{D_i}{h} \right) k_h(D_i) (\theta_t(D_i) - \mathbf{r}_p \left(\frac{D_i}{h} \right)^\top \eta_t(h)) \mathbf{1}_{\mathcal{J}_t}(D_i) \right] \right|. \end{aligned}$$

Assuming θ_t is s -times continuously differentiable, we have almost surely

$$\max_{1 \leq i \leq n} \left| \theta_t(D_i) - \mathbf{r}_p \left(\frac{D_i}{h} \right)^\top \eta_t(h) \right| \mathbf{1}(k_h(D_i)) \lesssim h^{\min\{s,p\}+1}.$$

By Lemma SA-2, $\|\boldsymbol{\Psi}_t^{-1}\| \lesssim 1$, and the same argument on each entry and finite dimensionality of the basis implies

$$\mathbb{E} \left[\left\| \mathbf{r}_p \left(\frac{D_i}{h} \right) k_h(D_i) \mathbf{1}_{\mathcal{J}_t}(D_i) \right\| \right] \lesssim 1.$$

The conclusion follows. $\square$

SA-3.5 Verification of High Level Conditions

This section verifies part (ii)-(iv) of Assumption SA-3, and the additional smoothness assumed in Lemma SA-7, for the special case when $\mathcal{B}$ is a finite collection of connected linear segments.

Lemma SA-8 (Piecewise Linear Boundary). *Suppose Assumption SA-2 holds with $d = 2$, and $\mathcal{d}(\cdot, \cdot)$ is the Euclidean distance. If $\mathcal{B}$ is piecewise linear with finite many pieces, then*

- (i) $J_1 d(\cdot, \mathcal{B}) = 1$ $\mathbf{m}$ -almost everywhere on $\mathcal{R}_t(\epsilon)$, and $\int_{\mathcal{R}_t(\epsilon)} \left| \frac{f(\mathbf{x})}{J_1 d(\mathbf{x}, \mathcal{B})} \right| d\mathbf{x} < \infty$,
- (ii) $\lim_{\epsilon \downarrow 0} F(s) = F(0)$ is finite, where $F(s) = \int_{d(\mathbf{x}, \mathcal{B})=s} \frac{f(\mathbf{x})}{J_1 d(\mathbf{x}, \mathcal{B})} d\mathfrak{H}^1(\mathbf{x})$ for all $s \geq 0$, and
- (iii) θ_t is continuous at 0,

for small enough $\epsilon > 0$, and $t \in \{0, 1\}$.

If, in addition, μ_0 , μ_1 and f are s -times continuously differentiable on $\mathcal{X}$, then

- (iv) θ_0 and θ_1 are s -times continuously differentiable on $[0, \epsilon]$.

Proof. First, we verify the Jacobian condition (i). Since $\mathcal{B}$ is piecewise linear, we denote the linear pieces by $\mathcal{B}_1, \dots, \mathcal{B}_M$ with $\mathcal{B} = \sqcup_{j=1}^M \mathcal{B}_j$. For small enough ϵ , for any $\mathbf{x} \in \mathcal{R}_t(\epsilon)$, the distance to the boundary depends on at most two linear pieces. That is, there exists $1 \leq j_1, j_2 \leq M$ such that

$$d(\mathbf{x}, \mathcal{B}) = d(\mathbf{x}, \mathcal{B}_{j_1} \cup \mathcal{B}_{j_2}),$$

as in Figure SA-1. For any point $\mathbf{x}$ on line segment $\overline{ce}$, the directional derivative of $d(\cdot, \mathcal{B})$ in the direction normal to $\overline{ce}$ is 1, and the directional derivative of $d(\cdot, \mathcal{B})$ in the direction parallel to $\overline{ce}$ is 0. Hence with $\nabla d(\cdot, \mathcal{B})|_{\mathbf{x}} = \mathbf{O}[1, 0]^\top$, where $\mathbf{O}$ is some orthonormal change of coordinate matrix. This shows $J_1 d(\cdot, \mathcal{B})|_{\mathbf{x}} = 1$. The same argument applies for $\mathbf{x}$ on the line segment $\overline{df}$ and on the segment $\overline{vp}$, $\overline{vq}$.

For any $\mathbf{x}$ on the arc $\overline{cd}$, the directional derivative of $d(\cdot, \mathcal{B})$ at $\mathbf{x}$ is 1 in the direction $\overrightarrow{O\mathbf{x}}$, and the directional derivative of $d(\cdot, \mathcal{B})$ at $\mathbf{x}$ is 1 in the direction orthogonal to $\overrightarrow{O\mathbf{x}}$. Thus $\nabla d(\cdot, \mathcal{B})|_{\mathbf{x}} = \mathbf{O}[1, 0]^\top$, where $\mathbf{O}$ is some orthonormal change of coordinate matrix, and hence $J_1 d(\cdot, \mathcal{B})|_{\mathbf{x}} = 1$.

This shows $J_1 d(\cdot, \mathcal{B}) = 1$ $\mathbf{m}$ -almost everywhere on $\mathcal{R}_t(\epsilon)$, and $\int_{\mathcal{R}_t(\epsilon)} \left| \frac{f(\mathbf{x})}{J_1 d(\mathbf{x}, \mathcal{B})} \right| d\mathbf{x} < \infty$ follows from continuity of f on compact support $\mathcal{X}$.

Next, we want to show continuity of F at 0 from (ii). It suffices to show $|F(s) - F(0)| \leq C\epsilon$ for all $s \in (0, \epsilon)$. Again, since $\mathcal{B}$ has finitely many pieces, with ϵ small enough, we can localize to Figure SA-1, and suppose $\mathcal{B} = \mathcal{B}_i \sqcup \mathcal{B}_j$. Consider $t = 0$, and the level set $\{\mathbf{x} \in \mathcal{A}_0 : d(\mathbf{x}, \mathcal{B}) = r\} = \overline{ce} \sqcup \overline{cd} \sqcup \overline{df}$. Define the projection on $\mathcal{B}$ by

$$\mathbf{P}(\mathbf{x}) = \arg \min_{\mathbf{y} \in \mathcal{B}} \|\mathbf{x} - \mathbf{y}\|.$$

Then

$$\begin{aligned} F(r) - F(0) &= \int_{\overline{ce} \sqcup \overline{df}} f(\mathbf{x}) d\mathfrak{H}^{d-1}(\mathbf{x}) + \int_{\overline{cd}} f(\mathbf{x}) d\mathfrak{H}^{d-1}(\mathbf{x}) - \int_{\mathcal{B}_i \sqcup \mathcal{B}_j} f(\mathbf{x}) d\mathfrak{H}^{d-1}(\mathbf{x}) \\ &= \int_{\overline{ce} \sqcup \overline{df}} (f(\mathbf{x}) - f(\mathbf{P}(\mathbf{x}))) d\mathfrak{H}^{d-1}(\mathbf{x}) + \int_{\overline{cd}} f(\mathbf{x}) d\mathfrak{H}^{d-1}(\mathbf{x}). \end{aligned}$$

The definition of $\mathbf{P}$ implies $\|\mathbf{x} - \mathbf{P}(\mathbf{x})\| \leq \epsilon$. Moreover, $\mathfrak{H}^{d-1}(\overline{cd}) \leq 2\pi\epsilon$. Since f is continuous on compact support $\mathcal{X}$, f is also uniform continuous on $\mathcal{X}$. Moreover, since $\mathcal{X}$ is compact, the pieces

$$\begin{cases} (L_i - s, -r), & s \in [0, L_i], \\ r(\cos(\frac{3}{2}\pi - s/r), \sin(3/2\pi - s/r)), & s \in [L_i, L_i + \theta r], \\ (s - L_i - \theta r)(\cos(\alpha), \sin(\alpha)) + r(\sin(\alpha), \cos(\alpha)), & s \in [L_i + \theta r, L_i + L_j + \theta r]. \end{cases}$$

Hence for any function $\psi : \mathbb{R}^2 \rightarrow \mathbb{R}$, we have

$$\begin{aligned} & \int_{\{\mathbf{x} \in \mathcal{A}_0 : d(\mathbf{x}, \mathcal{B}) = r\}} \psi(\mathbf{x}) d\mathfrak{H}^1(\mathbf{x}) \\ &= \int_0^{L_i + L_j + \theta r} \psi(\gamma_{r,0}(s)) \|\gamma'_{r,t}(s)\| ds \\ &= \int_0^{L_i + L_j + \theta r} \psi(\gamma_{r,0}(s)) ds \\ &= I_1(r) + I_2(r) + I_3(r), \end{aligned}$$

where

$$I_1(r) = \int_0^{L_i} \psi((L_i - s, -r)) ds,$$

and using a change variable with $u = s/r$,

$$\begin{aligned} I_2(r) &= \int_{L_i}^{L_i + \theta r} \psi(r(\cos(\frac{3}{2}\pi - s/r), \sin(3/2\pi - s/r))) ds \\ &= \int_{L_i}^{L_i + \theta} \psi(r(\cos(\frac{3}{2}\pi - u), \sin(3/2\pi - u))) r du, \end{aligned}$$

and using a change of variable with $v = s - L_i - \theta r$,

$$\begin{aligned} I_3(r) &= \int_{L_i + \theta r}^{L_i + L_j + \theta r} \psi((s - L_i - \theta r)(\cos(\alpha), \sin(\alpha)) + r(\sin(\alpha), \cos(\alpha))) ds \\ &= \int_0^{L_j} \psi(v(\cos(\alpha), \sin(\alpha)) + r(\sin(\alpha), \cos(\alpha))) dv. \end{aligned}$$

It follows that if ψ is s times continuously differentiable, then $r \mapsto \int_{\{\mathbf{x} \in \mathcal{A}_0 : d(\mathbf{x}, \mathcal{B}) = r\}} \psi(\mathbf{x}) d\mathfrak{H}^1(\mathbf{x})$ is also s times continuously differentiable.

Similarly, for $t = 1$, a curve length parametrization is given by

$$\begin{aligned} \gamma_{r,1}(s) &= \\ &\begin{cases} (L_j - s, s), & s \in [0, L_j - r \cot(\frac{\alpha}{2})], \\ (r \cot(\alpha/2), s) + (s - L_j + r \cot(\alpha/2))(\cos(\alpha), \sin(\alpha)), & s \in [L_j - r \cot(\frac{\alpha}{2}), L_i + L_j - 2r \cot(\frac{\alpha}{2})]. \end{cases} \end{aligned}$$

The smoothness of $L_j - r \cot(\frac{\alpha}{2})$ and $L_i + L_j - 2r \cot(\frac{\alpha}{2})$ in terms of r , smoothness of $\gamma_{r,1}(s)$ in r on each piece implies that for any function $\psi : \mathbb{R}^2 \rightarrow \mathbb{R}$ that is s times continuously differentiable,

we have $r \mapsto \int_{\{\mathbf{x} \in \mathcal{A}_1: d(\mathbf{x}, \mathcal{B})=r\}} \psi(\mathbf{x}) d\mathfrak{H}^1(\mathbf{x})$ is also s times continuously differentiable.

Hence if both μ_t and f are s times continuously differentiable, then both

$$r \rightarrow \int_{\{\mathbf{x} \in A_t: d(\mathbf{x}, \mathcal{B})=r\}} f(\mathbf{x}) d\mathfrak{H}^{d-1}(\mathbf{x})$$

and

$$r \rightarrow \int_{\{\mathbf{x} \in A_t: d(\mathbf{x}, \mathcal{B})=r\}} f(\mathbf{x}) \mu_t(\mathbf{x}) d\mathfrak{H}^{d-1}(\mathbf{x})$$

are s times continuously differentiable. Since we have assumed f continuous and positive on $\mathcal{X}$, it is bounded from below on $\mathcal{X}$. Hence

$$\theta_t(r) = \frac{\int_{\{\mathbf{x} \in A_t: d(\mathbf{x}, \mathcal{B})=r\}} f(\mathbf{x}) \mu_t(\mathbf{x}) d\mathfrak{H}^{d-1}(\mathbf{x})}{\int_{\{\mathbf{x} \in A_t: d(\mathbf{x}, \mathcal{B})=r\}} f(\mathbf{x}) d\mathfrak{H}^{d-1}(\mathbf{x})}$$

is also s times continuously differentiable, for $t = 0, 1$. $\square$

Corollary SA-2 (Bias for Piecewise Linear Boundary). *Suppose Assumption SA-2 holds with $d = 2$, and $\mathcal{d}(\cdot, \cdot)$ is the Euclidean distance. If $\mathcal{B}$ is piecewise linear with finite many pieces, and f is $(p+1)$ -times continuously differentiable, then*

$$\mathfrak{B}_{n,t} \lesssim h^{p+1},$$

for $t \in \{0, 1\}$.

Proof. The conclusion follows from Lemma SA-7 and Lemma SA-8. $\square$

SA-4 Post-Aggregation Approach

Under the assumptions imposed, for $t \in \{0, 1\}$, we have

$$\hat{\beta}_t(\mathbf{x}) = \mathbf{H}^{-1} \hat{\Gamma}_{t,\mathbf{x}}^{-1} \mathbb{E}_n \left[\mathbf{R}_p \left(\frac{\mathbf{X}_i - \mathbf{x}}{h} \right) K_h(\mathbf{X}_i - \mathbf{x}) Y_i \mathbf{1}(\mathbf{X}_i \in \mathcal{A}_t) \right],$$

where $\mathbf{H} = \text{diag}((h^{|\mathbf{v}|})_{0 \leq |\mathbf{v}| \leq p})$ with $\mathbf{v}$ running through all $\frac{d+p}{d!}$ multi-indices such that $|\mathbf{v}| \leq p$, and

$$\hat{\Gamma}_{t,\mathbf{x}} = \mathbb{E}_n \left[\mathbf{R}_p \left(\frac{\mathbf{X}_i - \mathbf{x}}{h} \right) \mathbf{R}_p \left(\frac{\mathbf{X}_i - \mathbf{x}}{h} \right)^\top K_h(\mathbf{X}_i - \mathbf{x}) \mathbf{1}(\mathbf{X}_i \in \mathcal{A}_t) \right].$$

The following assumption is the natural generalization of Assumption 3 in the paper to the case $d \geq 2$.

Assumption SA-4 (Kernel Function and Bandwidth). *Let $t \in \{0, 1\}$.*

- (i) $K : \mathbb{R}^d \rightarrow [0, \infty)$ is compact supported and Lipschitz continuous, or $K(\mathbf{u}) = \mathbf{1}(\mathbf{u} \in [-1, 1]^d)$.

(ii) $\liminf_{h \downarrow 0} \inf_{\mathbf{x} \in \mathcal{B}} \int_{\mathcal{A}_t} K_h(\mathbf{u} - \mathbf{x}) d\mathbf{u} \gtrsim 1$.

For $\mathbf{x}, \mathbf{x}_1, \mathbf{x}_2 \in \mathcal{B}$ and $t \in \{0, 1\}$, we introduce the following quantities:

$$\begin{aligned}\mathbf{\Gamma}_{t,\mathbf{x}} &= \mathbb{E} \left[\mathbf{R}_p \left(\frac{\mathbf{X}_i - \mathbf{x}}{h} \right) \mathbf{R}_p \left(\frac{\mathbf{X}_i - \mathbf{x}}{h} \right)^\top K_h(\mathbf{X}_i - \mathbf{x}) \mathbf{1}(\mathbf{X}_i \in \mathcal{A}_t) \right], \\ \widehat{\mathbf{\Sigma}}_{t,\mathbf{x}_1,\mathbf{x}_2} &= h^d \mathbb{E}_n \left[\mathbf{R}_p \left(\frac{\mathbf{X}_i - \mathbf{x}_1}{h} \right) \mathbf{R}_p \left(\frac{\mathbf{X}_i - \mathbf{x}_2}{h} \right)^\top K_h(\mathbf{X}_i - \mathbf{x}_1) K_h(\mathbf{X}_i - \mathbf{x}_2) \widehat{\varepsilon}_i(\mathbf{x}_1) \widehat{\varepsilon}_i(\mathbf{x}_2) \mathbf{1}(\mathbf{X}_i \in \mathcal{A}_t) \right], \\ \mathbf{\Sigma}_{t,\mathbf{x}_1,\mathbf{x}_2} &= h^d \mathbb{E} \left[\mathbf{R}_p \left(\frac{\mathbf{X}_i - \mathbf{x}_1}{h} \right) \mathbf{R}_p \left(\frac{\mathbf{X}_i - \mathbf{x}_2}{h} \right)^\top K_h(\mathbf{X}_i - \mathbf{x}_1) K_h(\mathbf{X}_i - \mathbf{x}_2) \sigma_t^2(\mathbf{X}_i) \mathbf{1}(\mathbf{X}_i \in \mathcal{A}_t) \right], \\ \widehat{\Omega}_{t,\mathbf{x}_1,\mathbf{x}_2} &= \frac{1}{nh^d} \mathbf{e}_1^\top \widehat{\mathbf{\Gamma}}_{t,\mathbf{x}_1}^{-1} \widehat{\mathbf{\Sigma}}_{t,\mathbf{x}_1,\mathbf{x}_2} \widehat{\mathbf{\Gamma}}_{t,\mathbf{x}_2}^{-1} \mathbf{e}_1, \quad \widehat{\Omega}_{\mathbf{x}_1,\mathbf{x}_2} = \widehat{\Omega}_{0,\mathbf{x}_1,\mathbf{x}_2} + \widehat{\Omega}_{1,\mathbf{x}_1,\mathbf{x}_2}, \\ \Omega_{t,\mathbf{x}_1,\mathbf{x}_2} &= \frac{1}{nh^d} \mathbf{e}_1^\top \mathbf{\Gamma}_{t,\mathbf{x}_1}^{-1} \mathbf{\Sigma}_{t,\mathbf{x}_1,\mathbf{x}_2} \mathbf{\Gamma}_{t,\mathbf{x}_2}^{-1} \mathbf{e}_1, \quad \Omega_{\mathbf{x}_1,\mathbf{x}_2} = \Omega_{0,\mathbf{x}_1,\mathbf{x}_2} + \Omega_{1,\mathbf{x}_1,\mathbf{x}_2},\end{aligned}$$

where $\widehat{\varepsilon}_i(\mathbf{x}) = Y_i - \sum_{t \in \{0,1\}} \mathbf{1}(\mathbf{X}_i \in \mathcal{A}_t) \widehat{\beta}_t(\mathbf{x})^\top \mathbf{R}_p(\mathbf{X}_i - \mathbf{x})$ and $\sigma_t^2(\mathbf{x}) = \mathbb{V}[Y_i(t) | \mathbf{X}_i = \mathbf{x}]$.

In addition, we have

$$\begin{aligned}\bar{B}_{t,\mathbf{x}} &= \mathbf{e}_1^\top \widehat{\mathbf{\Gamma}}_{t,\mathbf{x}}^{-1} \sum_{|\boldsymbol{\omega}|=p+1} \frac{\mu_t^{(\boldsymbol{\omega})}(\mathbf{x})}{\boldsymbol{\omega}!} \mathbb{E}_n \left[\mathbf{R}_p \left(\frac{\mathbf{X}_i - \mathbf{x}}{h} \right) \left(\frac{\mathbf{X}_i - \mathbf{x}}{h} \right)^\boldsymbol{\omega} K_h(\mathbf{X}_i - \mathbf{x}) \right], \quad \bar{B}_{\mathbf{x}} = \bar{B}_{1,\mathbf{x}} - \bar{B}_{0,\mathbf{x}}, \\ B_{t,\mathbf{x}} &= \mathbf{e}_1^\top \mathbf{\Gamma}_{t,\mathbf{x}}^{-1} \sum_{|\boldsymbol{\omega}|=p+1} \frac{\mu_t^{(\boldsymbol{\omega})}(\mathbf{x})}{\boldsymbol{\omega}!} \mathbb{E} \left[\mathbf{R}_p \left(\frac{\mathbf{X}_i - \mathbf{x}}{h} \right) \left(\frac{\mathbf{X}_i - \mathbf{x}}{h} \right)^\boldsymbol{\omega} K_h(\mathbf{X}_i - \mathbf{x}) \right], \quad B_{\mathbf{x}} = B_{1,\mathbf{x}} - B_{0,\mathbf{x}}, \\ \mathbf{Q}_{t,\mathbf{x}} &= \mathbb{E}_n \left[\mathbf{R}_p \left(\frac{\mathbf{X}_i - \mathbf{x}}{h} \right) K_h(\mathbf{X}_i - \mathbf{x}) \mathbf{1}(\mathbf{X}_i \in \mathcal{A}_t) u_i \right],\end{aligned}$$

where $u_i = Y_i - \sum_{t \in \{0,1\}} \mathbf{1}(\mathbf{X}_i \in \mathcal{A}_t) \mu_t(\mathbf{X}_i)$, and

$$\begin{aligned}\widehat{V}_{t,\mathbf{x}} &= h^{-1} \mathbf{e}_1^\top \widehat{\mathbf{\Gamma}}_{t,\mathbf{x}}^{-1} \widehat{\mathbf{\Sigma}}_{t,\mathbf{x},\mathbf{x}} \widehat{\mathbf{\Gamma}}_{t,\mathbf{x}}^{-1} \mathbf{e}_1, \quad \widehat{V}_{\mathbf{x}} = \widehat{V}_{0,\mathbf{x}} + \widehat{V}_{1,\mathbf{x}}, \\ V_{t,\mathbf{x}} &= h^{-1} \mathbf{e}_1^\top \mathbf{\Gamma}_{t,\mathbf{x}}^{-1} \mathbf{\Sigma}_{t,\mathbf{x},\mathbf{x}} \mathbf{\Gamma}_{t,\mathbf{x}}^{-1} \mathbf{e}_1, \quad V_{\mathbf{x}} = V_{0,\mathbf{x}} + V_{1,\mathbf{x}},\end{aligned}$$

Recall that we assumed, only for simplicity, that $\int_{\mathcal{B}} w(\mathbf{b}) dH^1(\mathbf{b}) = 1$. Then, the post-aggregation estimator becomes

$$\widehat{\tau}_w = \int_{\mathcal{B}} \widehat{\tau}(\mathbf{b}) w(\mathbf{b}) d\mathfrak{H}^{d-1}(\mathbf{b}).$$

Finally, we define the aggregated bias and variance quantities:

$$\begin{aligned}B_{\mathcal{B}} &= B_{1,\mathcal{B}} - B_{0,\mathcal{B}}, \quad B_{t,\mathcal{B}} = \int_{\mathcal{B}} B_{t,\mathbf{b}} w(\mathbf{b}) d\mathfrak{H}^{d-1}(\mathbf{b}), \\ \bar{B}_{\mathcal{B}} &= \bar{B}_{1,\mathcal{B}} - \bar{B}_{0,\mathcal{B}}, \quad \bar{B}_{t,\mathcal{B}} = \int_{\mathcal{B}} \bar{B}_{t,\mathbf{b}} w(\mathbf{b}) d\mathfrak{H}^{d-1}(\mathbf{b}), \\ \Omega_{\mathcal{B}} &= \Omega_{1,\mathcal{B}} + \Omega_{0,\mathcal{B}}, \quad \Omega_{t,\mathcal{B}} = \int_{\mathcal{B}} \int_{\mathcal{B}} \Omega_{t,\mathbf{b}_1,\mathbf{b}_2} w(\mathbf{b}_1) w(\mathbf{b}_2) d\mathfrak{H}^{d-1}(\mathbf{b}_1) d\mathfrak{H}^{d-1}(\mathbf{b}_2),\end{aligned}$$

$$\widehat{\Omega}_{\mathcal{B}} = \widehat{\Omega}_{1,\mathcal{B}} + \widehat{\Omega}_{0,\mathcal{B}}, \quad \widehat{\Omega}_{t,\mathcal{B}} = \int_{\mathcal{B}} \int_{\mathcal{B}} \widehat{\Omega}_{t,\mathbf{b}_1,\mathbf{b}_2} w(\mathbf{b}_1) w(\mathbf{b}_2) d\mathfrak{H}^{d-1}(\mathbf{b}_1) d\mathfrak{H}^{d-1}(\mathbf{b}_2),$$

for $t \in \{0, 1\}$.

SA-4.1 Preliminary Lemmas

This section states the preliminary results on matrix convergence. In what follows, we denote $\mathbf{X} = (\mathbf{X}_1^\top, \dots, \mathbf{X}_n^\top)$ and $\mathbf{W}_n = ((\mathbf{X}_1^\top, Y_1), \dots, (\mathbf{X}_n^\top, Y_n))^\top$.

Lemma SA-9 (Gram). *Suppose Assumption SA-2(i)-(ii) and Assumption SA-4 hold. If $\frac{\log(1/h)}{nh^d} = o(1)$, then*

$$\begin{aligned} \sup_{\mathbf{x} \in \mathcal{B}} \left\| \widehat{\Gamma}_{t,\mathbf{x}} - \Gamma_{t,\mathbf{x}} \right\| &\lesssim_{\mathbb{P}} \sqrt{\frac{\log(1/h)}{nh^d}}, \quad 1 \lesssim_{\mathbb{P}} \inf_{\mathbf{x} \in \mathcal{B}} \left\| \widehat{\Gamma}_{t,\mathbf{x}} \right\| \lesssim \sup_{\mathbf{x} \in \mathcal{B}} \left\| \widehat{\Gamma}_{t,\mathbf{x}} \right\| \lesssim_{\mathbb{P}} 1, \\ \sup_{\mathbf{x} \in \mathcal{B}} \left\| \widehat{\Gamma}_{t,\mathbf{x}}^{-1} - \Gamma_{t,\mathbf{x}}^{-1} \right\| &\lesssim_{\mathbb{P}} \sqrt{\frac{\log(1/h)}{nh^d}}, \end{aligned}$$

for $t \in \{0, 1\}$.

Proof. See Lemma 2.1 in the supplemental appendix of Cattaneo et al. [2025]. $\square$

Lemma SA-10 (Bias). *Suppose Assumption SA-2(i)-(iii) and Assumption SA-4 hold. If $\frac{\log(1/h)}{nh^d} = o(1)$, then*

$$\sup_{\mathbf{x} \in \mathcal{B}} |\mathbb{E}[\widehat{\mu}_t(\mathbf{x})|\mathbf{X}] - \mu_t(\mathbf{x})| \lesssim_{\mathbb{P}} h^{p+1},$$

for $t \in \{0, 1\}$. If, in addition, $h = o(1)$, then

$$\sup_{\mathbf{x} \in \mathcal{B}} |\mathbb{E}[\widehat{\mu}_t(\mathbf{x})|\mathbf{X}] - \mu_t(\mathbf{x}) - h^{p+1}\bar{B}_{t,\mathbf{x}}| = o_{\mathbb{P}}(h^{p+1}),$$

for $t \in \{0, 1\}$. Moreover, $\sup_{\mathbf{x} \in \mathcal{B}} |\bar{B}_{t,\mathbf{x}} - B_{t,\mathbf{x}}| \lesssim_{\mathbb{P}} \sqrt{\frac{\log(1/h)}{nh^d}}$, which implies $\sup_{\mathbf{x} \in \mathcal{B}} |\bar{B}_{t,\mathbf{x}}| \lesssim_{\mathbb{P}} 1$ for $t \in \{0, 1\}$.

Proof. See Lemma 2.2 in the supplemental appendix of Cattaneo et al. [2025]. $\square$

Lemma SA-11 (Aggregated Bias Along $\mathcal{B}$). *Suppose Assumption SA-2(i)-(iii) and Assumption SA-4 hold. If $\frac{\log(1/h)}{nh^d} = o(1)$ and $h = o(1)$, then*

$$\begin{aligned} \int_{\mathcal{B}} (\mathbb{E}[\widehat{\mu}_t(\mathbf{x})|\mathbf{X}] - \mu_t(\mathbf{x})) w(\mathbf{x}) d\mathfrak{H}^{d-1}(\mathbf{x}) &= h^{p+1} B_{\mathcal{B}} + o_{\mathbb{P}}(h^{p+1}) \\ &= h^{p+1} \bar{B}_{\mathcal{B}} + o_{\mathbb{P}}(h^{p+1}). \end{aligned}$$

Proof. The conclusion follows from Lemma SA-10 and the assumption that $\int_{\mathcal{B}} w(\mathbf{x}) d\mathfrak{H}^{d-1}(\mathbf{x}) = 1$. $\square$

Lemma SA-12 (Stochastic Linear Approximation). *Suppose Assumption SA-2(i)-(v) and Assumption SA-4 hold. If $\frac{\log(1/h)}{nh^d} = o(1)$, then*

$$\sup_{x \in \mathcal{B}} \|\mathbf{Q}_{t,\mathbf{x}}\| \lesssim \mathbb{P} \sqrt{\frac{\log(1/h)}{nh^d}} + \frac{\log(1/h)}{n^{\frac{1+v}{2+v}} h^d},$$

$$\sup_{\mathbf{x} \in \mathcal{B}} \left| \hat{\mu}_t(\mathbf{x}) - \mathbb{E}[\hat{\mu}_t(\mathbf{x})|\mathbf{X}] - \mathbf{e}_{1+\nu}^\top \mathbf{H}^{-1} \mathbf{\Gamma}_{t,\mathbf{x}}^{-1} \mathbf{Q}_{t,\mathbf{x}} \right| \lesssim \mathbb{P} \sqrt{\frac{\log(1/h)}{nh^d}} \left(\sqrt{\frac{\log(1/h)}{nh^d}} + \frac{\log(1/h)}{n^{\frac{1+v}{2+v}} h^d} \right),$$

for $t \in \{0, 1\}$.

Proof. See Lemma 2.3 in the supplemental appendix of Cattaneo et al. [2025]. $\square$

Lemma SA-13 (Covariance). *Suppose Assumptions SA-2 and SA-4 hold. If $\frac{\log(1/h)}{nh^d} = o(1)$, then*

$$\sup_{\mathbf{x}_1, \mathbf{x}_2 \in \mathcal{B}} \left\| \hat{\Sigma}_{t,\mathbf{x}_1, \mathbf{x}_2} - \Sigma_{t,\mathbf{x}_1, \mathbf{x}_2} \right\| \lesssim \mathbb{P} \sqrt{\frac{\log(1/h)}{nh^d}} + \frac{\log(1/h)}{n^{\frac{v}{2+v}} h^d} + h^{p+1},$$

$$\sup_{\mathbf{x}_1, \mathbf{x}_2 \in \mathcal{B}} \left| \hat{\Omega}_{\mathbf{x}_1, \mathbf{x}_2} - \Omega_{\mathbf{x}_1, \mathbf{x}_2} \right| \lesssim \mathbb{P} (nh^d)^{-1} \left(\sqrt{\frac{\log(1/h)}{nh^d}} + \frac{\log(1/h)}{n^{\frac{v}{2+v}} h^d} + h^{p+1} \right),$$

for $t \in \{0, 1\}$.

Proof. See Lemma 2.4 from the supplemental appendix of Cattaneo et al. [2025]. $\square$

Lemma SA-14 (Variance of Aggregated Estimator). *Suppose Assumptions SA-2 and SA-4 hold. If $\frac{\log(1/h)}{nh^d} = o(1)$ and $h = o(1)$, then*

$$\mathbb{V}[\hat{\tau}_w|\mathbf{X}] = \Omega_{\mathcal{B}} + O_{\mathbb{P}} \left(h^{d-1} \frac{\log(1/h)^{1/2}}{(nh^d)^{3/2}} \right) = \Omega_{\mathcal{B}} + o_{\mathbb{P}}((nh)^{-1}),$$

where

$$(nh)^{-1} \lesssim \Omega_{\mathcal{B}} \lesssim (nh)^{-1}.$$

If, in addition, $\frac{\log(1/h)}{n^{\frac{v}{2+v}} h^d} = o(1)$, then

$$\mathbb{V}[\hat{\tau}_w|\mathbf{X}] = \hat{\Omega}_{\mathcal{B}} + o_{\mathbb{P}}((nh)^{-1}).$$

Proof. Observe that $\mathbb{V}[\hat{\tau}_w|\mathbf{X}] = \sum_{t=0,1} \mathbb{V}[\int_{\mathcal{B}} \hat{\mu}_t(\mathbf{b}) w(\mathbf{b}) d\mathfrak{H}^{d-1}(\mathbf{b})]$, where for $t = 0, 1$,

$$\begin{aligned} & \mathbb{V} \left[\int_{\mathcal{B}} \hat{\mu}_t(\mathbf{b}) w(\mathbf{b}) d\mathfrak{H}^{d-1}(\mathbf{b}) \middle| \mathbf{X} \right] \\ &= \int_{\mathcal{B}} \int_{\mathcal{B}} \text{Cov} \left[\hat{\mu}_t(\mathbf{b}_1), \hat{\mu}_t(\mathbf{b}_2) \middle| \mathbf{X} \right] w(\mathbf{b}_1) w(\mathbf{b}_2) d\mathfrak{H}^{d-1}(\mathbf{b}_1) d\mathfrak{H}^{d-1}(\mathbf{b}_2). \end{aligned}$$

Consider

$$\bar{\Sigma}_{t,\mathbf{b}_1,\mathbf{b}_2} = h^d \mathbb{E}_n \left[\mathbf{R}_p \left(\frac{\mathbf{X}_i - \mathbf{b}_1}{h} \right) \mathbf{R}_p \left(\frac{\mathbf{X}_i - \mathbf{b}_2}{h} \right)^\top K_h(\mathbf{X}_i - \mathbf{b}_1) K_h(\mathbf{X}_i - \mathbf{b}_2) \sigma_t^2(\mathbf{X}_i) \mathbf{1}(\mathbf{X}_i \in \mathcal{A}_t) \right]. \quad (\text{SA-7})$$

Then the same argument as the proof of Lemma SA-9 implies if $\frac{\log(1/h)}{nh^d} = o(1)$, then

$$\sup_{\mathbf{b}_1, \mathbf{b}_2 \in \mathcal{B}} \|\bar{\Sigma}_{t,\mathbf{b}_1,\mathbf{b}_2} - \Sigma_{t,\mathbf{b}_1,\mathbf{b}_2}\| = O_{\mathbb{P}} \left(\sqrt{\frac{\log(1/h)}{nh^d}} \right). \quad (\text{SA-8})$$

Together with Lemma SA-9, we have when $\frac{\log(1/h)}{nh^d} = o(1)$,

$$\begin{aligned} & \sup_{\mathbf{b}_1, \mathbf{b}_2 \in \mathcal{B}} \|\text{Cov}[\hat{\mu}_t(\mathbf{b}_1), \hat{\mu}_t(\mathbf{b}_2) | \mathbf{X}] - \Omega_{t,\mathbf{b}_1,\mathbf{b}_2}\| \\ &= \sup_{\mathbf{b}_1, \mathbf{b}_2 \in \mathcal{B}} \|(nh^d)^{-1} \mathbf{e}_1^\top \hat{\Gamma}_t^{-1} \bar{\Sigma}_{t,\mathbf{b}_1,\mathbf{b}_2} \hat{\Gamma}_t^{-1} \mathbf{e}_1 - \Omega_{t,\mathbf{b}_1,\mathbf{b}_2}\| \\ &= O_{\mathbb{P}} \left(\frac{\log(1/h)^{1/2}}{(nh^d)^{3/2}} \right). \end{aligned}$$

We have assumed that K is supported on a compact set. W.l.o.g, assume the support of K has a diameter no greater than R , $0 < R < \infty$. Consider the set

$$\mathcal{E}(h) = \{(\mathbf{x}, \mathbf{y}) \in \mathcal{B} \times \mathcal{B} : \|\mathbf{x} - \mathbf{y}\| \leq hR\}. \quad (\text{SA-9})$$

Since $\mathcal{B}$ is $d-1$ dimensional, we have $\mathbf{m}(\mathcal{E}(h)) \lesssim h^{d-1}$, where $\mathbf{m}$ is the $2(d-1)$ dimensional Lebesgue measure. Hence

$$\begin{aligned} & \mathbb{V} \left[\int_{\mathcal{B}} \hat{\mu}_t(\mathbf{b}) w(\mathbf{b}) d\mathfrak{H}^{d-1}(\mathbf{b}) | \mathbf{X} \right] - \Omega_t \\ &= \int_{\mathcal{B}} \int_{\mathcal{B}} \left(\text{Cov}[\hat{\mu}_t(\mathbf{b}_1), \hat{\mu}_t(\mathbf{b}_2) | \mathbf{X}] - \Omega_{t,\mathbf{b}_1,\mathbf{b}_2} \right) w(\mathbf{b}_1) w(\mathbf{b}_2) d\mathfrak{H}^{d-1}(\mathbf{b}_1) d\mathfrak{H}^{d-1}(\mathbf{b}_2) \\ &\lesssim \sup_{\mathbf{b}_1, \mathbf{b}_2 \in \mathcal{B}} \|\text{Cov}[\hat{\mu}_t(\mathbf{b}_1), \hat{\mu}_t(\mathbf{b}_2) | \mathbf{X}] - \Omega_{t,\mathbf{b}_1,\mathbf{b}_2}\| \cdot \\ &\quad \int_{\mathcal{B}} \int_{\mathcal{B}} \mathbf{1}((\mathbf{b}_1, \mathbf{b}_2) \in \mathcal{E}(h)) w(\mathbf{b}_1) w(\mathbf{b}_2) d\mathfrak{H}^{d-1}(\mathbf{b}_1) d\mathfrak{H}^{d-1}(\mathbf{b}_2) \\ &= O_{\mathbb{P}} \left(h^{d-1} \frac{\log(1/h)^{1/2}}{(nh^d)^{3/2}} \right) \\ &= o_{\mathbb{P}}((nh)^{-1}), \end{aligned} \quad (\text{SA-10})$$

where in the last line we have used the assumption that $\frac{\log(1/h)}{nh^d} = o(1)$. This proves the first claim.

For the second argument,

$$\Omega_{t,\mathcal{B}} = \int_{\mathcal{B}} \int_{\mathcal{B}} \Omega_{t,\mathbf{b}_1,\mathbf{b}_2} w(\mathbf{b}_1) w(\mathbf{b}_2) d\mathfrak{H}^{d-1}(\mathbf{b}_1) d\mathfrak{H}^{d-1}(\mathbf{b}_2)$$

$$\begin{aligned}
&\leq \sup_{\mathbf{b}_1, \mathbf{b}_2 \in \mathcal{B}} |\Omega_{t, \mathbf{b}_1, \mathbf{b}_2}| \int_{\mathcal{B}} \int_{\mathcal{B}} \mathbf{1}((\mathbf{b}_1, \mathbf{b}_2) \in \mathcal{E}(h)) w(\mathbf{b}_1) w(\mathbf{b}_2) d\mathfrak{H}^{d-1}(\mathbf{b}_1) d\mathfrak{H}^{d-1}(\mathbf{b}_2) \\
&\lesssim (nh^d)^{-1} \mathbf{m}(\mathcal{E}(h)) \\
&\lesssim (nh)^{-1}.
\end{aligned}$$

This shows the upper bound. For the lower bound, let $\mathbf{b}_1 \in \mathcal{B}$ and $\mathbf{b}_2 = \mathbf{b}_1 + h\delta$ for some vector delta such that $\sup_{\mathbf{x} \in \mathcal{X}} K_h(\mathbf{x} - \mathbf{b}_1) K_h(\mathbf{x} - \mathbf{b}_2) > 0$. A change of variable then implies a typical element of $\Sigma_{t, \mathbf{b}_1, \mathbf{b}_2}$ has the form

$$\begin{aligned}
&\mathbb{E} \left[\left(\frac{\mathbf{X}_i - \mathbf{b}_1}{h} \right)^{\mathbf{u}} \left(\frac{\mathbf{X}_i - \mathbf{b}_1 - \delta h}{h} \right)^{\mathbf{v}} \frac{1}{h^d} K \left(\frac{\mathbf{X}_i - \mathbf{b}_1}{h} \right) K \left(\frac{\mathbf{X}_i - \mathbf{b}_1 - \delta h}{h} \right) \sigma_t^2(\mathbf{X}_i) \mathbf{1}(\mathbf{X}_i \in \mathcal{A}_t) \right] \\
&= \int_{\mathbf{b}_1 + h\mathcal{A}_t} \mathbf{s}^{\mathbf{u}} (\mathbf{s} - \delta)^{\mathbf{v}} K(\mathbf{s}) K(\mathbf{s} + \delta) \sigma_t^2(\mathbf{b}_1 + h\mathbf{s}) d\mathbf{s} \\
&\gtrsim 1.
\end{aligned}$$

It follows that $|\Omega_{t, \mathbf{b}_1, \mathbf{b}_2}| \gtrsim (nh^d)^{-1}$ for $(\mathbf{b}_1, \mathbf{b}_2)$ on a set $\mathcal{E}'(h)$ such that $\mathbf{m}(\mathcal{E}'(h)) \gtrsim h^{d-1}$. This leads to the lower bound in the second claim.

The third claim follows from Lemma SA-13 and the same analysis as Equation (SA-10). $\square$

SA-4.2 Mean Square Error Expansion

Theorem SA-3 (MSE Expansions). *Suppose Assumptions SA-2 and SA-4 hold. If $\frac{\log(1/h)}{n^{\frac{v}{2+v}} h^d} = o(1)$ and $h = o(1)$, then*

$$\mathbb{E}[(\hat{\tau}_w - \tau_w)^2 | \mathbf{X}] = \Omega_{\mathcal{B}} + h^{2p+2} B_{\mathcal{B}}^2 + o_{\mathbb{P}}((nh)^{-1}) + o_{\mathbb{P}}(h^{2p+2}).$$

Proof. Using the decomposition

$$\mathbb{E}[(\hat{\tau}_w - \tau_w)^2 | \mathbf{X}] = \mathbb{V}[\hat{\tau}_w | \mathbf{X}] + (\mathbb{E}[\hat{\tau}_w | \mathbf{X}] - \tau_w)^2,$$

the conclusion then follows from Lemmas SA-11 and SA-14. $\square$

SA-4.3 Central Limit Theorem

The feasible t -statistics is

$$\hat{T}_w = \frac{\hat{\tau}_w - \tau_w}{\sqrt{\hat{\Omega}_{\mathcal{B}}}}.$$

Theorem SA-4 (Asymptotic Normality). *Suppose Assumptions SA-2 and SA-4 hold. If $\frac{\log(1/h)}{n^{\frac{v}{2+v}} h^d} =$*

$o(1)$ and $(nh)h^{2p+2} = o(1)$, then

$$\sup_{u \in \mathbb{R}} |\mathbb{P}(\hat{\mathbf{T}}_w \leq u) - \Phi(u)| = o(1).$$

Proof. We decompose the error $\hat{\tau}_w - \tau_w = (\hat{\mu}_{1,w} - \mu_{1,w}) - (\hat{\mu}_{0,w} - \mu_{0,w})$, where

$$\begin{aligned} \hat{\mu}_{t,w} - \mu_{t,w} &= \int_{\mathcal{B}} (\hat{\mu}_1(\mathbf{x}) - \mu_1(\mathbf{x})) w(\mathbf{x}) d\mathfrak{H}^{d-1}(\mathbf{b}) \\ &= \underbrace{\int_{\mathcal{B}} \mathbf{e}_1^\top \mathbf{\Gamma}_{t,\mathbf{b}}^{-1} \mathbf{Q}_{t,\mathbf{b}} d\mathfrak{H}^{d-1}(\mathbf{b})}_{\text{linear error}} + \underbrace{\int_{\mathcal{B}} \mathbf{e}_1^\top (\hat{\mathbf{\Gamma}}_{t,\mathbf{b}}^{-1} - \mathbf{\Gamma}_{t,\mathbf{b}}^{-1}) \mathbf{Q}_{t,\mathbf{b}} d\mathfrak{H}^{d-1}(\mathbf{b})}_{\text{non-linear error}} + O_{\mathbb{P}}(h^{p+1}) \end{aligned}$$

for $t \in \{0, 1\}$, and using Lemma SA-11 to bound the approximation error

For the non-linearity error, using $\bar{\mathbf{\Sigma}}_{t,\mathbf{b}_1,\mathbf{b}_2}$ in Equation (SA-7),

$$\begin{aligned} &\mathbb{E} \left[\left(\int_{\mathcal{B}} \mathbf{e}_1^\top (\hat{\mathbf{\Gamma}}_{t,\mathbf{b}}^{-1} - \mathbf{\Gamma}_{t,\mathbf{b}}^{-1}) \mathbf{Q}_{t,\mathbf{b}} d\mathfrak{H}^{d-1}(\mathbf{b}) \right)^2 \middle| \mathbf{X} \right] \\ &= \int_{\mathcal{B}} \int_{\mathcal{B}} \mathbf{e}_1^\top (\hat{\mathbf{\Gamma}}_{t,\mathbf{b}_1}^{-1} - \mathbf{\Gamma}_{t,\mathbf{b}_1}^{-1}) (nh^d)^{-1} \bar{\mathbf{\Sigma}}_{t,\mathbf{b}_1,\mathbf{b}_2} (\hat{\mathbf{\Gamma}}_{t,\mathbf{b}_2}^{-1} - \mathbf{\Gamma}_{t,\mathbf{b}_2}^{-1}) \mathbf{e}_1 w(\mathbf{b}_1) w(\mathbf{b}_2) d\mathfrak{H}^{d-1}(\mathbf{b}_1) d\mathfrak{H}^{d-1}(\mathbf{b}_2). \end{aligned}$$

Since $\bar{\mathbf{\Sigma}}_{t,\mathbf{b}_1,\mathbf{b}_2} = 0$ if $\mathbf{b}_1$ and $\mathbf{b}_2$ are farther away from each other than the diameter of $\text{Supp}(K)$, we can use Equation (SA-9) to get

$$\begin{aligned} &\int_{\mathcal{B}} \int_{\mathcal{B}} \mathbf{e}_1^\top (\hat{\mathbf{\Gamma}}_{t,\mathbf{b}_1}^{-1} - \mathbf{\Gamma}_{t,\mathbf{b}_1}^{-1}) (nh^d)^{-1} \bar{\mathbf{\Sigma}}_{t,\mathbf{b}_1,\mathbf{b}_2} (\hat{\mathbf{\Gamma}}_{t,\mathbf{b}_2}^{-1} - \mathbf{\Gamma}_{t,\mathbf{b}_2}^{-1}) \mathbf{e}_1 d\mathfrak{H}^{d-1}(\mathbf{b}_1) d\mathfrak{H}^{d-1}(\mathbf{b}_2) \\ &\leq \sup_{\mathbf{b} \in \mathcal{B}} (\hat{\mathbf{\Gamma}}_{t,\mathbf{b}}^{-1} - \mathbf{\Gamma}_{t,\mathbf{b}}^{-1})^2 \sup_{\mathbf{b}_1, \mathbf{b}_2 \in \mathcal{B}} \|\bar{\mathbf{\Sigma}}_{t,\mathbf{b}_1,\mathbf{b}_2}\| \sup_{\mathbf{b} \in \mathcal{B}} |w(\mathbf{b})| (nh^d)^{-1} \mathfrak{m}(\mathcal{Z}(h)) \\ &\lesssim \sup_{\mathbf{b} \in \mathcal{B}} (\hat{\mathbf{\Gamma}}_{t,\mathbf{b}}^{-1} - \mathbf{\Gamma}_{t,\mathbf{b}}^{-1})^2 \sup_{\mathbf{b}_1, \mathbf{b}_2 \in \mathcal{B}} \|\bar{\mathbf{\Sigma}}_{t,\mathbf{b}_1,\mathbf{b}_2}\| (nh)^{-1}, \end{aligned}$$

where the constant does not depend on n or $\mathbf{X}$. Then, by Lemma SA-9 and Equation (SA-8),

$$\begin{aligned} &\int_{\mathcal{B}} \mathbf{e}_1^\top (\hat{\mathbf{\Gamma}}_{t,\mathbf{b}}^{-1} - \mathbf{\Gamma}_{t,\mathbf{b}}^{-1}) \mathbf{Q}_{t,\mathbf{b}} d\mathfrak{H}^{d-1}(\mathbf{b}) \\ &\lesssim_{\mathbb{P}} \left[\sup_{\mathbf{b} \in \mathcal{B}} (\hat{\mathbf{\Gamma}}_{t,\mathbf{b}}^{-1} - \mathbf{\Gamma}_{t,\mathbf{b}}^{-1})^2 \sup_{\mathbf{b}_1, \mathbf{b}_2 \in \mathcal{B}} \|\bar{\mathbf{\Sigma}}_{t,\mathbf{b}_1,\mathbf{b}_2}\| (nh)^{-1} \right]^{1/2} \\ &\lesssim_{\mathbb{P}} o_{\mathbb{P}}((nh)^{-1/2}) \end{aligned}$$

because $\frac{\log(1/h)}{nh^d} = o(1)$.

Consider now the stochastic linearized T-statistic

$$\bar{\mathbf{T}}_w = \Omega_{\mathcal{B}}^{-1/2} \int_{\mathcal{B}} \mathbf{e}_1^\top \mathbf{\Gamma}_{t,\mathbf{b}}^{-1} \mathbf{Q}_{t,\mathbf{b}} d\mathfrak{H}^{d-1}(\mathbf{b}).$$

Then, by Lemma SA-14 and the previous bounds,

$$\widehat{\mathbf{T}}_w - \overline{\mathbf{T}}_w = (\widehat{\Omega}_{\mathcal{B}}^{-1/2} - \Omega_{\mathcal{B}}^{-1/2}) \int_{\mathcal{B}} \mathbf{e}_1^\top \Gamma_{t,\mathbf{b}}^{-1} \mathbf{Q}_{t,\mathbf{b}} d\mathfrak{H}^{d-1}(\mathbf{b}) + o_{\mathbb{P}}(1) = o_{\mathbb{P}}(1) \quad (\text{SA-11})$$

because $\frac{\log(1/h)}{n^{\frac{p}{2+p}} h^d} = o(1)$ and $(nh)h^{2p+2} = o(1)$.

Finally, we apply the Berry-Esseen lemma to the linearized statistic $\overline{\mathbf{T}}_w = \sum_{i=1}^n Z_i$, where

$$Z_i = n^{-1} \Omega_{\mathcal{B}}^{-1/2} \int_{\mathcal{B}} \mathbf{e}_1^\top \Gamma_{t,\mathbf{b}}^{-1} \mathbf{R}_p \left(\frac{\mathbf{X}_i - \mathbf{b}}{h} \right) K_h(\mathbf{X}_i - \mathbf{b}) \mathbf{1}(\mathbf{X}_i \in \mathcal{A}_t) u_i w(\mathbf{b}) d\mathfrak{H}^{d-1}(\mathbf{b}),$$

which satisfies $\mathbb{E}[Z_i] = 0$. The definition of $\Omega_{\mathcal{B}}$ implies that $\sum_{i=1}^n \mathbb{V}[Z_i] = \Omega_{\mathcal{B}}^{-1/2} \Omega_{\mathcal{B}} \Omega_{\mathcal{B}}^{-1/2} = 1$. Hence,

$$\begin{aligned} & \sum_{i=1}^n \mathbb{E}[|Z_i^3|] \\ &= n^{-3} \Omega_{\mathcal{B}}^{-3/2} \sum_{i=1}^n \mathbb{E} \left[\left(\int_{\mathcal{B}} \mathbf{e}_1^\top \Gamma_{t,\mathbf{b}}^{-1} \mathbf{R}_p \left(\frac{\mathbf{X}_i - \mathbf{b}}{h} \right) K_h(\mathbf{X}_i - \mathbf{b}) \mathbf{1}(\mathbf{X}_i \in \mathcal{A}_t) u_i w(\mathbf{b}) d\mathfrak{H}^{d-1}(\mathbf{b}) \right)^3 \right]. \end{aligned}$$

Consider

$$G(\mathbf{b}_1, \mathbf{b}_2, \mathbf{b}_3) = g(\mathbf{X}_i, u_i, \mathbf{b}_1) g(\mathbf{X}_i, u_i, \mathbf{b}_2) g(\mathbf{X}_i, u_i, \mathbf{b}_3),$$

where

$$g(\mathbf{X}_i, u_i, \mathbf{b}) = \mathbf{e}_1^\top \Gamma_{t,\mathbf{b}}^{-1} \mathbf{R}_p \left(\frac{\mathbf{X}_i - \mathbf{b}}{h} \right) K_h(\mathbf{X}_i - \mathbf{b}) \mathbf{1}(\mathbf{X}_i \in \mathcal{A}_t) u_i.$$

The same argument as Lemma SA-9 shows that

$$\sup_{\mathbf{b}_1, \mathbf{b}_2, \mathbf{b}_3 \in \mathcal{B}} \mathbb{E}[|G(\mathbf{b}_1, \mathbf{b}_2, \mathbf{b}_3)|] \lesssim h^{-2d}$$

when $\frac{\log(1/n)}{nh^d} = o(1)$. Suppose the support of K has diameter less than R , and consider

$$\mathcal{F}(h) = \{(\mathbf{b}_1, \mathbf{b}_2, \mathbf{b}_3) \in \mathcal{B}^3 : \|\mathbf{b}_i - \mathbf{b}_j\| \leq R, j = 1, 2, 3\}.$$

Since $\mathcal{B}$ is $d-1$ dimensional, $\mathbf{m}(\mathcal{F}(h)) \lesssim h^{2d-2}$. It follows that

$$\begin{aligned} & \mathbb{E} \left[\left(\int_{\mathcal{B}} \mathbf{e}_1^\top \Gamma_{t,\mathbf{b}}^{-1} \mathbf{R}_p \left(\frac{\mathbf{X}_i - \mathbf{b}}{h} \right) K_h(\mathbf{X}_i - \mathbf{b}) \mathbf{1}(\mathbf{X}_i \in \mathcal{A}_t) u_i w(\mathbf{b}) d\mathfrak{H}^{d-1}(\mathbf{b}) \right)^3 \right] \\ &= \mathbb{E} \left[\int_{\mathbf{b}_1 \in \mathcal{B}} \int_{\mathbf{b}_2 \in \mathcal{B}} \int_{\mathbf{b}_3 \in \mathcal{B}} G(\mathbf{b}_1, \mathbf{b}_2, \mathbf{b}_3) w(\mathbf{b}_1) w(\mathbf{b}_2) w(\mathbf{b}_3) d\mathfrak{H}^{d-1}(\mathbf{b}_1) d\mathfrak{H}^{d-1}(\mathbf{b}_2) d\mathfrak{H}^{d-1}(\mathbf{b}_3) \right] \\ &\lesssim \mathbf{m}(\mathcal{F}(h)) \sup_{\mathbf{b}_1, \mathbf{b}_2, \mathbf{b}_3 \in \mathcal{B}} \mathbb{E}[|G(\mathbf{b}_1, \mathbf{b}_2, \mathbf{b}_3)|] \lesssim nh^{-2}. \end{aligned}$$

Together with the rate of $\Omega_{\mathcal{B}}$ from Lemma [SA-14](#), we have

$$\sum_{i=1}^n \mathbb{E}[|Z_i^3|] \lesssim (nh)^{-1/2}.$$

By Berry-Esseen lemma, we have

$$\sup_{u \in \mathbb{R}} |\mathbb{P}(\bar{T}_w \leq u) - \Phi(u)| = O((nh)^{-1/2}),$$

and the final conclusion then follows from Equation ([SA-11](#)). □

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